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Highlights

- Recovery with multiple prior information via solving n-l1 minimization is proposed
- Theoretical measurement bounds required by the n-l1 minimization are established
- The derived bounds of the n-l1 minimization are sharper
- The proposed n-l1 minimization outperforms the state-of-the-art algorithms

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Sparse Signal Recovery with Multiple Prior Information: Algorithm and Measurement Bounds $^{\bigstar}$

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Abstract

We address the problem of reconstructing a sparse signal from compressive measurements with the aid of multiple known correlated signals. We propose a reconstruction algorithm with multiple side information signals (RAMSI), which solves an n- ℓ_1 minimization problem by weighting adaptively the multiple side information signals at every iteration. In addition, we establish theoretical bounds on the number of measurements required to guarantee successful reconstruction of the sparse signal via weighted n- ℓ_1 minimization. The analysis of the derived bounds reveals that weighted n- ℓ_1 minimization can achieve sharper bounds and significant performance improvements compared to classical compressed sensing (CS). We evaluate experimentally the proposed RAMSI algorithm and the established bounds using numerical sparse signals. The results show that the proposed algorithm outperforms state-of-the-art algorithms—including classical CS, ℓ_1 - ℓ_1 minimization, Modified-CS, regularized Modified-CS, and weighted ℓ_1 minimization—in terms of both the theoretical bounds and the practical performance.

Keywords: Compressed sensing, prior information, weighted $n-\ell_1$ minimization, measurement bounds

1. Introduction

Compressed sensing (CS) [1–15] states that sparse signals can be recovered in a computationally tractable manner from a limited set of measurements by minimizing the ℓ_1 -norm. The CS performance can be improved by replacing the ℓ_1 -norm with a weighted ℓ_1 -norm [8, 9, 16–18]. The studies in [11, 12] provide bounds on the number of measurements required for successful signal recovery based on convex

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optimization. Furthermore, distributed compressed sensing [13, 14] allows a correlated ensemble of sparse signals to be jointly recovered by exploiting the intra- and inter-signal correlations.

We consider the problem of reconstructing a signal given side or prior information, gleaned from a set of known correlated signals. Initially, this problem was studied in [16, 19–28], where the modified-CS method [19, 21] considered that a part of the support is available from prior knowledge and tried

- ¹⁰ CS method [19, 21] considered that a part of the support is available from prior knowledge and tried to find the signal that satisfies the measurement constraint and is sparsest outside the known support. Prior information on the sparsity pattern of the data was also considered in [23] and informationtheoretic guarantees were presented. The studies in [24, 25] introduced weights into the ℓ_1 minimization framework that depend on the partitioning of the source signal into two sets, with the entries of each set having a specific probability of being nonzero.
- Alternatively, the studies in [29–32] incorporated side information on CS by means of ℓ_1 - ℓ_1 minimization and derived bounds on the number of Gaussian measurements required to guarantee perfect signal recovery. It was shown that ℓ_1 - ℓ_1 minimization can dramatically improve the reconstruction performance over CS subject to a good-quality side information [29, 30]. The study in [33] proposed a weighted- ℓ_1 minimization method to incorporate prior information—in the form of inaccurate support estimates—into CS. The work also provided bounds on the number of Gaussian measurements for successful recovery when sufficient support information is available. Furthermore, recent studies proposed CS with side information in practical applications, i.e., compressive video foreground extraction [34–36], magnetic resonance imaging (MRI) [17], and synthetic aperture radar imaging [37]. The problem of sparse signal recovery with prior information also emerges in the context of recon-
- structing a sequence of time-varying sparse signals from low-dimensional measurements [23, 38–42]. The study in [42] reviewed a class of recursive algorithms for recovering a sequence of time-varying sparse signals from a small number of measurements. The problem of estimating a time-varying, sparse signal from streaming measurements was studied in [38, 42], while the work in [23] addressed the recovery problem in the context of multiple measurement vectors. The problem also appears in the context of robust PCA and online robust PCA [39–41], a framework that finds important application in video
- background subtraction. The studies in [39–41] used the modified-CS [21, 22] method to leverage prior knowledge under the condition of slow support and signal value changes.

1.1. Motivation

Emerging applications require reconstructing a sequence of signals from compressive measurements; for example, dynamic MRI [43] involves the acquisition of moving cardiovascular, abdominal, or larynx images. Related research questions are also encountered in the context of compressive video sensing [44, 45]. Alternative applications [46–48] follow a distributed sensing scenario, where a plethora of tiny heterogeneous devices collect information from the environment, and require sensing and processing

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- ⁴⁰ under time and resource constraints. These challenges can be addressed by leveraging the distributed sparse representation of the multiple sources [48–50]. The problem in this setup is to reconstruct the sparse sources along with exploiting the correlation among them; alias, a key question is how to reconstruct a signal from a small number of measurements, by leveraging efficiently the redundancy gleaned from multiple correlated signals.
- ⁴⁵ Existing attempts to incorporate prior information in compressed sensing are typically considering one known prior information signal of good quality; see for example [17, 29, 30, 34, 35]. Conversely, we are aiming at reconstructing a sparse signal from a small set of measurements with the aid of *multiple prior signals*, which change in time; that is, there are arbitrary prior information qualities and varying correlations among them. This raises key interesting questions:
- How can we leverage efficiently the correlation across multiple prior information signals? This implies a strategy to exploit the useful information from the multiple signals and alleviate negative effects due to low quality prior information.
 - How many measurements are required to successfully reconstruct the sparse signal given multiple prior signals? This calls for bounds on the number of measurements required to guarantee successful signal recovery.

1.2. Contributions

To address these questions, we contribute in a twofold way. Firstly, we propose a novel sparse signal Reconstruction Algorithm that leverages Multiple Side Information (RAMSI). Secondly, we establish lower bounds on the number of measurements required by RAMSI to recover the sparse signal.

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The RAMSI algorithm solves a *(re-)weighted n-l*₁ minimization problem: Per iteration of the recovery process, the algorithm weights adaptively the *n* prior information signals. As such, unlike existing works [17, 29, 30, 34, 35], which exploit only one prior information signal, RAMSI can efficiently leverage the correlations among multiple signals and adapt on-the-fly to changes of the correlations. We show experimentally that RAMSI leads to higher recovery performance than state-of-the-art methods [7, 21, 22, 25, 29].

We also establish measurement bounds for the weighted $n-\ell_1$ minimization problem, which serve as lower bounds for the RAMSI algorithm. The bounds depend on the support of the source signal to be recovered and the correlations between the target signal and the multiple prior information signals. The correlations are expressed via the supports of the differences between the source and prior information signals. We will show that the weighted $n-\ell_1$ minimization bounds are sharper compared to those of the classical CS [1, 3, 5] and the $\ell_1-\ell_1$ reconstruction [29, 30] methods. These bounds

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depict the advantage of RAMSI to deal with heterogeneous side information signals including possible

poor prior information signals. Furthermore, we show—both theoretically and practically—that the performance of the method is improved with the number of available prior information signals.

75 1.3. Outline

The rest of this paper is as follows: Section 2 reviews the background on CS and CS with side information. The RAMSI algorithm is proposed in Section 3, whereas our measurement bounds and their analysis are presented in Section 4. The bounds and the performance of RAMSI are assessed in Section 5 and Section 6 concludes the work.

80 2. Background

2.1. Compressed Sensing

Let $x \in \mathbb{R}^n$ be a signal that is sparse in a domain or learned set of basis. The signal can be reduced to a vector $y = \Phi x \in \mathbb{R}^m$ by sampling with a measurement matrix $\Phi \in \mathbb{R}^{m \times n}$ $(m \ll n)$ whose elements are sampled from an i.i.d. Gaussian distribution. The signal can be recovered by solving the Basis Pursuit problem [3, 5, 52]:

$$\min_{\boldsymbol{x}} \|\boldsymbol{x}\|_1 \text{ subject to } \boldsymbol{y} = \boldsymbol{\Phi} \boldsymbol{x}, \tag{1}$$

where $\|\boldsymbol{x}\|_1 := \sum_{i=1}^n |x_i|$ is the ℓ_1 -norm of \boldsymbol{x} wherein x_i is an element of \boldsymbol{x} . Problem (1) becomes an instance of finding a general solution:

$$\min_{\boldsymbol{x}} \{ H(\boldsymbol{x}) = f(\boldsymbol{x}) + g(\boldsymbol{x}) \},$$
(2)

where $f := \mathbb{R}^n \to \mathbb{R}$ is a smooth convex function and $g := \mathbb{R}^n \to \mathbb{R}$ is a continuous convex function, possibly non-smooth. Specifically, in Problem (1) we have $g(\boldsymbol{x}) = \lambda \|\boldsymbol{x}\|_1$ and $f(\boldsymbol{x}) = \frac{1}{2} \|\boldsymbol{\Phi}\boldsymbol{x} - \boldsymbol{y}\|_2^2$, where $\|\cdot\|_2$ denotes ℓ_2 -norm, with Lipschitz constant $L_{\nabla f}$ [7]. Using proximal gradient methods [7], $\boldsymbol{x}^{(k)}$ at iteration k is computed as

$$\boldsymbol{x}^{(k)} = \Gamma_{\frac{1}{L}g} \Big(\boldsymbol{x}^{(k-1)} - \frac{1}{L} \nabla f(\boldsymbol{x}^{(k-1)}) \Big),$$
(3)

where $L \ge L_{\nabla f}$, and $\Gamma_{\frac{1}{L}g}(\boldsymbol{x})$ is the proximal operator:

$$\Gamma_{\frac{1}{L}g}(\boldsymbol{x}) = \operatorname*{arg\,min}_{\boldsymbol{v} \in \mathbb{R}^n} \left\{ \frac{1}{L} g(\boldsymbol{v}) + \frac{1}{2} \|\boldsymbol{v} - \boldsymbol{x}\|_2^2 \right\}.$$
(4)

The classical ℓ_1 minimization problem in CS [1, 3, 5] requires m_{ℓ_1} measurements [11, 29, 30] for successful reconstruction, bounded as

$$m_{\ell_1} \ge 2s_0 \log \frac{n}{s_0} + \frac{7}{5}s_0 + 1,\tag{5}$$

where $s_0 := \|\boldsymbol{x}\|_0 = |\{i : x_i \neq 0\}|$ denotes the number of nonzero elements in \boldsymbol{x} as the support of \boldsymbol{x} , with |.| denoting the cardinality of a set and $\|\cdot\|_0$ being the ℓ_0 -pseudo-norm.

2.2. CS with Support Knowledge

The modified-CS method in [19, 21] considered prior knowledge in the form of knowing a part of the signal's support. Furthermore, the study in [33] proposed a weighted- ℓ_1 minimization method to incorporate prior support information and provided bounds on the number of measurements—we denote this number by $m_{\omega\ell_1}$ —for successful recovery. It is worth noting that the bound of the number of measurements required by modified-CS, m_{modCS} , is a special case of the bound on $m_{\omega\ell_1}$. Namely, without considering weights on the prior support, it holds that [33, Corrollary 8] $m_{\omega\ell_1} \equiv m_{\text{modCS}} \geq$ $s_0 + (1 + C^{-2})s_e \log(en/s_e)$, where C < 1 and s_e is the size of the support estimate error defined by $|(\tilde{T} \cap T^c) \cup (T \cap \tilde{T^c})| \leq s_e$, where T is the support, \tilde{T} is the support estimate and T^e and $\tilde{T^c}$ their respective complements. In [33], it is assumed that $s_0 \leq n/2$ and $s_e \leq s_0$. Due to C < 1, we can write

$$m_{\omega\ell_1} \equiv m_{\text{modCS}} \ge 2s_e \log(n/s_e) + 2s_e + s_0.$$
(6)

⁸⁵ 2.3. CS with ℓ_1 - ℓ_1 Minimization

The ℓ_1 - ℓ_1 minimization approach [29, 30, 34] reconstructs x given a signal $z \in \mathbb{R}^n$ as prior information by solving the following problem:

$$\min_{\boldsymbol{x}} \Big\{ \frac{1}{2} \| \boldsymbol{\Phi} \boldsymbol{x} - \boldsymbol{y} \|_{2}^{2} + \lambda (\| \boldsymbol{x} \|_{1} + \| \boldsymbol{x} - \boldsymbol{z} \|_{1}) \Big\}.$$
(7)

The bound on the number of measurements required by Problem (7) to successfully reconstruct \boldsymbol{x} depends on the quality of the side information signal \boldsymbol{z} as [29, 30, 34]

$$m_{\ell_1-\ell_1} \ge 2\overline{h}\log\left(\frac{n}{s_0+\xi/2}\right) + \frac{7}{5}\left(s_0+\frac{\xi}{2}\right) + 1,$$
(8)

where

$$\xi := |\{i : z_i \neq x_i = 0\}| - |\{i : z_i = x_i \neq 0\}|,$$
(9a)

$$\bar{h} := |\{i : x_i > 0, x_i > z_i\} \cup \{i : x_i < 0, x_i < z_i\}|,$$
(9b)

wherein x_i , z_i are corresponding elements of x, z. It has been shown that the solution of Problem (7) improves over the solution of Problem (1) provided that the prior information has good enough quality [29, 30]. The quality is expressed by a high number of elements z_i that are equal to x_i , thereby leading to ξ in (9a) being small.

⁹⁰ 3. Recovery With Multiple Prior Information

3.1. Problem Statement

We consider the problem of recovering x from a low-dimensional measurement vector y given prior information. Classical CS methods can be used to recover x from y; however, these methods do



Figure 1: Illustration of a compressive video sensing setup in (a), and the vectorized wavelet transform coefficients of (b) frame no. 1286, (c) frame no. 1285, and (d) frame no. 1284 in the Bootstrap video sequence. In the considered scenario, the representation of frame no. 1286 can be reconstructed using the recovered representations of the frames no. 1285 and 1284 as prior information.

not leverage the correlations across multiple different data vectors. Examples of practical applications that motivate our work include compressive video sensing and analysis [35, 51, 53] and dynamic MRI 95 reconstruction [45]. For instance, consider the compressive video sensing setup illustrated in Fig. 1(a): an incoming video frame is sensed using compressive measurements at the encoder and reconstructed at the decoder with the aid of multiple prior information, gleaned from at set of previously reconstructed frames. Fig. 1(b) depicts the vectorized wavelet transform coefficients (a.k.a. sparse representation) of frame no. 1286 in the Bootstrap video sequence—which corresponds to the target signal x in 100 our scenario and the wavelet representations of the two previous reconstructed frames¹ no. 1285 [Fig. 1(c)] and 1284 [Fig. 1(d)]—which respectively correspond to z_1 and z_2 in our scenario. It is clear that the sparse representation x is highly correlated with z_1 and z_2 ; this correlation can be leveraged to improve the reconstruction of x. Other application scenarios include compressive foreground extraction [35] and compressive online robust principle component analysis [53], where a foreground frame x can be reconstructed from previously reconstructed foregrounds z_1 and z_2 .

¹Alternatively, the prior information could be gleaned from motion-compensated versions of the reconstructed frames [54].

The $\ell_1 - \ell_1$ minimization framework in Problem (7) can be used to recover \boldsymbol{x} from \boldsymbol{y} given either \boldsymbol{z}_1 or \boldsymbol{z}_2 , that is, only one prior information vector. Moreover, it can be that $\ell_1 - \ell_1$ minimization performs worse than the ℓ_1 minimization method due to low correlation between the target signal and the prior information signal, e.g., \boldsymbol{z}_2 . To address these two limitations, we propose a new *reconstruction algorithm with multiple side information* (RAMSI). The input of RAMSI is the measurement vector $\boldsymbol{y} =$ $\boldsymbol{\Phi}\boldsymbol{x}$ and J prior information signals $\boldsymbol{z}_1, \ldots, \boldsymbol{z}_J \in \mathbb{R}^n$. The objective function is constructed by using an n- ℓ_1 -norm function in Problem (2):

$$g(\boldsymbol{x}) = \lambda \sum_{j=0}^{J} \| \mathbf{W}_{j}(\boldsymbol{x} - \boldsymbol{z}_{j}) \|_{1},$$
(10)

where $\mathbf{z}_0 = \mathbf{0}$ and \mathbf{W}_j are diagonal weight matrices, $\mathbf{W}_j = \text{diag}(w_{j1}, w_{j2}, ..., w_{jn})$, wherein $w_{ji} > 0$ is the weight in \mathbf{W}_j at index *i*. Namely, the objective function of the proposed n- ℓ_1 minimization problem is given by

$$\min_{\boldsymbol{x}} \Big\{ H(\boldsymbol{x}) = \frac{1}{2} \|\boldsymbol{\Phi}\boldsymbol{x} - \boldsymbol{y}\|_{2}^{2} + \lambda \sum_{j=0}^{J} \|\mathbf{W}_{j}(\boldsymbol{x} - \boldsymbol{z}_{j})\|_{1} \Big\}.$$
(11)

3.2. The Proposed RAMSI Algorithm

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We solve Problem (11) by iteratively computing \boldsymbol{x} and updating \mathbf{W}_j based on the proximal gradient method [7]. The question is then how to determine the weight values in order to improve the reconstruction by effectively leveraging the multiple side information signals. This also calls for a method that avoids recovery performance degradation when the correlation among the prior information signals and the target signal decreases. As such, unlike prior studies [29, 30, 34], our method distributes weights across *multiple side information signals*. We propose to solve the problem in (11) such that at every iteration k we iterate over (i) a weight update step (i.e., updating \mathbf{W}_j given \boldsymbol{x}) and (ii) a data computation step (i.e., computing \boldsymbol{x} given \mathbf{W}_j).

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Updating \mathbf{W}_j given x and z_j . Fixing x and z_j , $j = 1, \ldots, J$, we determine the weights $\{w_{ji}\}$ so as to improve the reconstruction of x, where j indexes the side information signal and $i = 1, \ldots, n$ iterates over the data elements. To normalize the contribution of each side information signal during the iterative process, we impose a constraint on the weights. We may have different strategies to update the weights $\{w_{ji}\}$; in this work, use the constraint $\sum_{j=0}^{J} \mathbf{W}_j = \mathbf{I}_n$, where \mathbf{I}_n is the identity matrix with dimensions $n \times n$. Consider the *i*-th element in x in Problem (11): we want to assign $w_{ji} > 0$ to the related term $x_i - z_{ji}$, where z_{ji} is the *i*-th element in z_j . We aim at assigning to each element $x_i - z_{ji}$ a weight w_{ji} that is high when $|x_i - z_{ji}|$ is low; in this way, we can favor x_i components that are closer to z_{ji} . To ensure that zero-valued elements $|x_i - z_{ji}|$ do not cause a breakdown of the method, we add a small parameter $\epsilon > 0$ and we set

$$w_{ji} = \frac{\eta_i}{|x_i - z_{ji}| + \epsilon},\tag{12}$$

where $\eta_i > 0$. Under the constraint $\sum_{j=0}^{J} \mathbf{W}_j = \mathbf{I}_n$, we obtain

$$\eta_i = \left(\sum_{j=0}^J \frac{1}{|x_i - z_{ji}| + \epsilon}\right)^{-1}.$$
(13)

Using (13), we can rewrite each weight w_{ji} as

$$w_{ji} = \frac{(|x_i - z_{ji}| + \epsilon)^{-1}}{\sum_{l=0}^{J} (|x_i - z_{li}| + \epsilon)^{-1}}.$$
(14)

Computing x given \mathbf{W}_j . Given \mathbf{W}_j , RAMSI computes $x^{(k)}$ at iteration k using (3), where the proximal operator $\Gamma_{\frac{1}{r},g}(x_i)$ is computed as follows.

Proposition 3.1. The proximal operator $\Gamma_{\frac{1}{L}g}(\boldsymbol{x})$ in (4) for the problem of signal recovery with multiple side information, for which $g(\boldsymbol{x}) = \lambda \sum_{j=0}^{J} \|\mathbf{W}_{j}(\boldsymbol{x} - \boldsymbol{z}_{j})\|_{1}$, is given by

$$\Gamma_{\frac{1}{L}g}(x_i) = \begin{cases} x_i - \frac{\lambda}{L} \sum_{j=0}^{J} w_{ji}(-1)^{\mathfrak{b}(r(15)$$

where

$$z_{ri} + \frac{\lambda}{L} \sum_{j=0}^{J} w_{ji}(-1)^{\mathfrak{b}(r(16a)$$

$$z_{ri} + \frac{\lambda}{L} \sum_{j=0}^{J} w_{ji}(-1)^{\mathfrak{b}(r-1(16b)$$

and where, without loss of generality, we have assumed that $-\infty = z_{(-1)i} \leq z_{0i} \leq z_{1i} \leq \ldots \leq z_{Ji} \leq z_{(J+1)i} = \infty$, and we have defined a boolean function

$$\mathbf{b}(r < j) = \begin{cases} 1, & \text{if } r < j \\ 0, & \text{otherwise.} \end{cases}$$
(17)

with $r \in \{-1, ..., J\}$.

Proof. The proof is given in Appendix B.

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We summarize RAMSI in Algorithm 1 (the Matlab code is provided in [55]), which is based on the fast iterative soft-thresholding algorithm (FISTA) algorithm [7]. The *Stopping criterion* (cf. Line 2 of Algorithm 1) can be a maximum iteration number k_{\max} , a relative variation of the objective function $H(\mathbf{x})$ in (11), or a change of the number of nonzero components of the estimate $\mathbf{x}^{(k)}$. In this work, the relative variation of $H(\mathbf{x})$ is chosen as stopping criterion.

¹²⁵ 4. Bounds For Weighted $n-\ell_1$ Minimization

We now establish measurement bounds for weighted $n-\ell_1$ minimization, serving as lower bounds for RAMSI.

Algorithm 1: The proposed RAMSI algorithm. Input: $y, \Phi, z_1, z_2, ..., z_J;$ Output: \widehat{x} ; 1 Initialization : $\mathbf{W}_{0}^{(1)} = \mathbf{I}$; $\mathbf{W}_{i}^{(1)} = \mathbf{0} \ \forall \ 1 \le j \le J$; $\boldsymbol{u}^{(1)} = \boldsymbol{x}^{(0)} = \mathbf{0}$; $L = L_{\nabla f}$; $\lambda > 0$; $\epsilon > 0$; $t_1 = 1; k = 0;$ 2 while Stopping criterion is false do k = k + 1;3 $\nabla f(\boldsymbol{u}^{(k)}) = \boldsymbol{\Phi}^{\mathrm{T}}(\boldsymbol{\Phi}\boldsymbol{u}^{(k)} - \boldsymbol{y});$ 4 $\boldsymbol{x}^{(k)} = \Gamma_{\frac{1}{L}g} \left(\boldsymbol{u}^{(k)} - \frac{1}{L} \nabla f(\boldsymbol{u}^{(k)}) \right); \text{ where } \Gamma_{\frac{1}{L}g}(.) \text{ is given by (15)};$ 5 // Updating weights. $w_{ji}^{(k+1)} = \frac{(|x_i^{(k)} - z_{ji}| + \epsilon)^{-1}}{\sum\limits_{i=1}^{J} (|x_i^{(k)} - z_{li}| + \epsilon)^{-1}};$ 6 // Updating values for next iteration. $t_{k+1} = (1 + \sqrt{1 + 4t_k^2})/2;$ 7 $u^{(k+1)} = x^{(k)} + \frac{t_k - 1}{t_{k+1}} (x^{(k)} - x^{(k-1)});$ 8 9 end 10 return $x^{(k)}$; 4.1. Measurement Bound

4.1.1. Signal Traits

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We begin our analysis by stating Definitions 4.1, 4.2, and 4.3, which do not limit the generality of our analysis but help us formalizing our bounds. Let s_j denote the support of each difference vector $\boldsymbol{x}-\boldsymbol{z}_j$; namely, $\|\boldsymbol{x}-\boldsymbol{z}_j\|_0 = s_j$, where $j \in \{0, \ldots, J\}$ and $\boldsymbol{z}_0 = \boldsymbol{0}$.

Definition 4.1. There are $0 \le p < n$ indices for which the values in all $\{x - z_j\}_{j=0}^J$ vectors are nonzero and there are n - q indices, with $0 < q \le n$, for which the values in all $\{x - z_j\}_{j=0}^J$ vectors are equal to zero.

By rearranging the elements of all difference vectors in the same ordering, we can write the difference vectors as

$$\boldsymbol{x} - \boldsymbol{z}_{0} = (x_{1}, ..., x_{p}, x_{p+1}, ..., x_{q}, 0, ..., 0),$$

$$\boldsymbol{x} - \boldsymbol{z}_{1} = (x_{1} - z_{11}, ..., x_{p} - z_{1p}, x_{p+1} - z_{1(p+1)}, ..., x_{q} - z_{1q}, 0, ..., 0),$$

$$...$$

$$\boldsymbol{x} - \boldsymbol{z}_{J} = (x_{1} - z_{J1}, ..., x_{p} - z_{Jp}, x_{p+1} - z_{J(p+1)}, ..., x_{q} - z_{Jq}, 0, ..., 0).$$

$$(18)$$

Consider the following example:

$$\begin{aligned} \boldsymbol{x} &= (3, \ 0, \ 0, -2, \ 0, \ 0, \ 4, \ 0, \ 0, \ 0, \ 7), \\ \boldsymbol{x} - \boldsymbol{z}_1 &= (0, \ 0, \ 3, \ 0, \ 0, \ 0, \ 2, \ 0, \ 0, -1, \ 0), \\ \boldsymbol{x} - \boldsymbol{z}_2 &= (4, \ 0, \ 6, \ 0, \ 0, \ 3, -1, \ 0, \ 0, \ 0, \ 8). \end{aligned}$$

In this example, n = 11, $s_0 = 4$, $s_1 = 3$, and $s_2 = 5$. Since at index i = 7 the elements of all $\{x - z_j\}_{j=0}^2$ vectors are nonzero, we have p = 1. Moreover, the elements of all vectors are zero at index i = 2, 5, 8, 9, hence q = 7. After rearranging the elements in the same ordering of the vectors as in (18), we obtain

$$\begin{aligned} \boldsymbol{x} &= (4, 3, 0, -2, 0, 0, 7, 0, 0, 0, 0), \\ \boldsymbol{x} - \boldsymbol{z}_1 &= (2, 0, 3, 0, 0, -1, 0, 0, 0, 0, 0), \\ \boldsymbol{x} - \boldsymbol{z}_2 &= (-1, 4, 6, 0, 3, 0, 8, 0, 0, 0, 0). \end{aligned}$$
(19)

It is evident that $p \leq \min\{s_j\}$ and $q \geq \max\{s_j\}$. As such, p and q can be seen as parameters that express the common part of the support and of the zero positions across the difference vectors. It can be noted that if q = 0, $\boldsymbol{x} = \boldsymbol{z}_j = \boldsymbol{0}$ and if p = n, $\|\boldsymbol{x}\|_0 = n$, where \boldsymbol{x} is not a sparse source as our consideration.

Definition 4.2. Following the rearrangement in (18), at any given position $i \in \{p + 1, ..., q\}$, we define d_i as the number of zero elements out of the J+1 elements $\{x_i - z_{ji}\}_{j=0}^J$. Under Definition 4.1, we have that $d_i \in \{1, ..., J\}$.

In the example in (19), at i = 3 we have one zero element across the vectors, thus $d_3 = 1$. At i = 5 there are two zero elements, hence, $d_5 = 2$.

Using Definitions 4.1 and 4.2, we can express the total number of zero elements in (18) as

$$(J+1)(n-q) + \sum_{i=p+1}^{q} d_i = (J+1)n - \sum_{j=0}^{J} s_j.$$
 (20)

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Definition 4.3. Let *i* be a given position at the signal and multiple side information vectors. For any index *i*, without loss of generality, rearrange the J + 1 elements $\{x_i - z_{ji}\}_{j=0}^J$ in order given by increasing-values order into $\{x_i - z_{j'i}\}_{j'=0}^J$, i.e., $-\infty = z_{(-1)i} \leq z_{0i} \leq z_{1i} \leq \ldots \leq z_{Ji} \leq z_{(J+1)i} = \infty$, where, for convenience, we introduced $z_{(-1)i}$ and $z_{(J+1)i}$ to denote $-\infty$ and ∞ , respectively. It can be noted that we rearrange for each individual index *i*. Under Definitions 4.1 and 4.2, we define $r_i \in \{-1, \ldots, J\}$ such that $x_i \in (z_{r_ii}, z_{(r_i+1)i}]$, or alias, $\operatorname{sign}(x_i - z_{j'i}) = (-1)^{\mathfrak{b}(r_i < j')}$, where $\mathfrak{b}(r_i < j')$ is defined in (17).

At the example in (19), at index i = 1, $x_1 = 4$ and $\{z_{j'1}\}_{j'=-1}^3 = \{-\infty, 0, 2, 5, \infty\}$. Thus $x_1 \in (z_{11}, z_{21}]$, i.e., $r_1 = 1$. For convenience, in the rest of the paper, we use notation j instead of j' denoting the reordered indices.

4.1.2. The Measurement Bound

Based on Definitions 4.1, 4.2, and 4.3, we shall derive a bound on the number of measurements required by weighted $n-\ell_1$ minimization to successfully recover the target signal. Our generic bound is defined by Theorem 4.4, while a simpler but looser bound is described in Section 4.2.

Theorem 4.4. The number of measurements $m_{n-\ell_1}$ required by weighted $n-\ell_1$ minimization to recover the signal \boldsymbol{x} , given measurements $\boldsymbol{y} = \boldsymbol{\Phi} \boldsymbol{x}$ and J side information signals \boldsymbol{z}_j , is bounded as

$$m_{n-\ell_1} \ge 2\bar{a}_{n-\ell_1} \log \frac{n}{\bar{s}_{n-\ell_1}} + \frac{7}{5}\bar{s}_{n-\ell_1} + \delta_{n-\ell_1} + 1,$$
(21)

where $\bar{a}_{n-\ell_1}$, $\bar{s}_{n-\ell_1}$, and $\delta_{n-\ell_1}$ are defined as

$$\bar{a}_{n-\ell_1} = \sum_{i=1}^{p} a_i^2,$$
 (22a)

$$\bar{s}_{n-\ell_1} = p + \sum_{i=p+1} (1-c_i),$$
(22b)

$$\delta_{n-\ell_1} = (\kappa_{n-\ell_1} - 1)(\bar{s}_{n-\ell_1} - p), \qquad (22c)$$

wherein
$$a_i = \sum_{j=0}^{J} w_{ji}(-1)^{\mathfrak{b}(r_i < j)}, c_i = d_i \left(\sum_{j=0}^{J} \frac{\epsilon}{|x_i - z_{ji}| + \epsilon}\right)^{-1}$$
 with $\epsilon > 0$, and

$$\kappa_{n \cdot \ell_1} = \frac{2 \cdot \min\{c_i\}}{\sqrt{\pi \log(n/\tilde{s}_{n \cdot \ell_1})}(2 \cdot \min\{c_i\} - 1)}.$$
(23)

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Proof. The proof is given in Appendix Appendix C.

The code computing the quantities is provided in [55].

4.2. Further Analysis and Comparison with Known Bounds

We now relate the derived bound in Theorem 4.4 with the bounds of compressed sensing and compressed sensing with prior information, which are reported in Sec. 2.

Corollary 4.4.1 (Relations with the known bounds). The bound $m_{n-\ell_1}$ for weighted $n-\ell_1$ minimization in (21) becomes

(a) the ℓ_1 -minimization bound m_{ℓ_1} in (5), when $\mathbf{W}_0 = \mathbf{I}_n$ and $\mathbf{W}_j = \mathbf{0}$ for $j \in \{1, \ldots, J\}$, that is,

$$m_{n-\ell_1} \equiv m_{\ell_1} \ge 2s_0 \log \frac{n}{s_0} + \frac{7}{5}s_0 + 1, \tag{24}$$

(b) the ℓ_1 - ℓ_1 minimization bound $m_{\ell_1-\ell_1}$ in (8), when $\mathbf{W}_0 = \mathbf{W}_1 = \frac{1}{2}\mathbf{I}_n$ and $\mathbf{W}_j = \mathbf{0}$ for $j \in \{2, \ldots, J\}$, that is,

$$m_{n-\ell_1} \equiv m_{\ell_1-\ell_1} \ge 2\bar{h}\log\frac{n}{\bar{s}_{\ell_1-\ell_1}} + \frac{7}{5}\bar{s}_{\ell_1-\ell_1} + 1,$$
(25)

where \bar{h} is given by (9b) and $\bar{s}_{\ell_1-\ell_1} = \frac{s_0+s_1}{2}$.

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Proof. The proof is given in Appendix Appendix C.

We now compare our bound for weighted $n-\ell_1$ minimization with the $\ell_1-\ell_1$ minimization [29, 30, 34] bound in (8). To this end, we derive a simplified bound in (26), which approximates our bound in (21). Furthermore, we introduce two simplicified bounds, which are looser bounds—one for each method—that are independent from the values of \boldsymbol{x} , \boldsymbol{z}_j .

The simplified bound. Our bound for weighted $n-\ell_1$ minimization in (21) becomes approximately

$$\widetilde{m}_{n-\ell_1} \ge 2\overline{a}_{n-\ell_1}\log\frac{n}{p} + \frac{7}{5}p + 1.$$
 (26)

Our approximation has the following reasoning: Firstly, Lemma Appendix D.1 in Appendix Appendix D proves that the $\delta_{n-\ell_1}$ term in (21) is negative; hence, the bound in (26) is looser. Secondly, since $\epsilon > 0$ is very small, we have $c_i \rightarrow 1^-$ [see Lemma Appendix D.1 for a quick explanation]. Consequently, $\bar{s}_{n-\ell_1} \approx p$ and $\delta_{n-\ell_1} \approx 0$ from (22b) and (22c), respectively; thereby, leading to our approximation. This simple bound is easier to evaluate compared to the bound in (21), as we only need to compute $\bar{a}_{n-\ell_1}$ and p. Furthermore, according to (22a) and Definition 4.1, we can write that $\bar{a}_{n-\ell_1} \leq p \leq \min\{s_j\}$, that is, p in (26) is smaller than s_0 in (5) and $\bar{s}_{\ell_1-\ell_1}$ in (25). Consequently, the simple bound in (26) for weighted $n-\ell_1$ minimization is sharper than the ℓ_1 minimization and the $\ell_1-\ell_1$ minimization bounds, which means that the bound in (21) is even sharper. For the special case that p = 0, $\bar{a}_{n-\ell_1} = 0$ from (22a), $\bar{s}_{n-\ell_1} \approx p$, $\delta_{n-\ell_1} \approx 0$, consequently, from (21) $\tilde{m}_{n-\ell_1} \geq 1$.

The simplified bounds with element-value independence. Weighted $n-\ell_1$ minimization and $\ell_1-\ell_1$ minimization have looser bounds that are independent from the values of \boldsymbol{x} , \boldsymbol{z}_j , as given by

$$\widehat{m}_{n-\ell_1} \ge 2p\log\frac{n}{p} + \frac{7}{5}p + 1,$$
(27)

$$\widehat{m}_{\ell_1 - \ell_1} \ge 2\rho \log \frac{n}{\bar{s}_{\ell_1 - \ell_1}} + \frac{7}{5} \bar{s}_{\ell_1 - \ell_1} + 1, \tag{28}$$

where p is defined in Definition 4.1, $\rho = \min\{s_0, s_1\}$, and $\bar{s}_{\ell_1-\ell_1} = \frac{s_0+s_1}{2}$. We obtain these bounds as follows: The bounds in (26) and (25) depend on the values of \boldsymbol{x} and \boldsymbol{z}_j , with $j \in \{1, \ldots, J\}$, via the quantities $\bar{a}_{n-\ell_1}$ and \bar{h} , respectively. From (22a) and (9b), we observe that $\bar{a}_{n-\ell_1} \leq p \leq \min\{s_j\}$ and $\bar{h} \leq \min\{s_0, s_1\}$. Hence, by replacing $\bar{a}_{n-\ell_1}$ and \bar{h} with their maximum value leads to the looser bounds in (27) and in (28).

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The bounds in (27) and (28) reveal the advantage of using multiple side information signals in compressed sensing: It is evident that, by definition, $p \leq \rho$; hence, $\hat{m}_{n-\ell_1} \leq \hat{m}_{\ell_1-\ell_1}$ showing that the $n-\ell_1$ bound is better than the $\ell_1-\ell_1$ bound. Moreover, the higher the number of side information signals the smaller the bound is expected to become (because $p \leq \min\{s_j\}$). 195

Furthermore, according to (28), if the side information signal z_1 is not good enough, i.e., if $s_1 \gg s_0$, then $\hat{m}_{\ell_1-\ell_1} > m_{\ell_1}$ because of $\bar{s}_{\ell_1-\ell_1} \gg s_0$, thereby highlighting the limitations of the $\ell_1-\ell_1$ minimization method compared to the proposed weighted $n-\ell_1$ minimization approach.

Finally, we compare our bound to the existing bound of the weighted ℓ_1 minimization with the prior support, $m_{\omega \ell_1}$ in (6). This bound depends on the support s_0 and the support estimate's error s_e from prior information that reveals that $m_{\omega \ell_1}$ in (6) is worse than $\widehat{m}_{n-\ell_1}$ in (26). If we do not have good support for the estimate, i.e., high values of s_e , then $m_{\omega \ell_1}$ in (6) is worse than m_{ℓ_1} in (5), which is illustrated in Sec. 5.

5. Experimental Results

5.1. Experimental Setup

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We consider the reconstruction of a synthetic sparse signal x given known prior signals z_i , j =1,2,3. We generate x from the i.i.d. zero-mean, unit-variance Gaussian distribution, with n = 1000 and $s_0 = 128$. Firstly, we consider the scenario where the side information signals z_j are highly correlated with the signal \boldsymbol{x} ; in this case, \boldsymbol{z}_j , j = 1, 2, 3, are generated such that they satisfy $s_j = \|\boldsymbol{x} - \boldsymbol{z}_j\|_0 = 64$ similar to the generation in [29, 30]. Moreover, a parameter is controlling the number of positions of nonzeros for which both x and $x - z_j$ coincide. Let us denote this number by ζ_j . For instance, if $\zeta_j = 51, x$ has 51 nonzero positions that coincide with 51 nonzero positions of $x - z_j$. This incurs a significant error between the source and the side information, given by $\|\boldsymbol{z}_j - \boldsymbol{x}\|_2 / \|\boldsymbol{x}\|_2 \approx 0.56$. We also generate prior information signals that are poorly correlated with x; in this case, we have $s_j = 256$ and $s_j = 352$. Furthermore, we set $\zeta_j = 128$, namely, 128 nonzero positions of x coincide with 128 nonzero positions out of the total 256 or 352 nonzero positions of the side information signals, z_j , j = 1, 2, 3. 215 This leads to very high errors, e.g., $\|\boldsymbol{z}_j - \boldsymbol{x}\|_2 / \|\boldsymbol{x}\|_2 \approx 1.12$ for $s_j = 256$, and the supports s_j of $\boldsymbol{x} - \boldsymbol{z}_j$ are much higher than that of x. To constrain the number of cases, we set all s_i equal.

5.2. Performance Evaluation

5.2.1. Signal Recovery Accuracy

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We now assess the performance of RAMSI against state-of-the-art methods. Furthermore, we evaluate the bounds for weighted $n-\ell_1$ minimization [cf. (21)] against the bounds for classical CS $[cf. (5)], \ell_1-\ell_1 \text{ minimization [cf. (8)], weighted}-\ell_1 \text{ and Modified-CS [cf. (6)]. For a fixed number$ of measurements m, the probability of successful recovery $\Pr(\text{success})$ is the number of times \boldsymbol{x} is recovered as \hat{x} with an error $\|\hat{x} - x\|_2 / \|x\|_2 \le 10^{-2}$, divided by the total number of 100 Monte-Carlo iterations (trials), where each trial considers different generated x, z_1, z_2, z_3, Φ .

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Let RAMSI-J- ℓ_1 denote the RAMSI algorithm that uses J side information signals, that is, one (\boldsymbol{z}_1) , two $(\boldsymbol{z}_1, \boldsymbol{z}_2)$, or three $(\boldsymbol{z}_1, \boldsymbol{z}_2, \boldsymbol{z}_3)$, where we set $\epsilon = 10^{-5}$, $\lambda = 10^{-5}$. The existing FISTA



Figure 2: Successful probabilities of the original 1000-D \boldsymbol{x} vs. number of measurements m for RAMSI using one, two, three side information signals.

[7], FISTA- ℓ_1 - ℓ_1 [30], Mod-CS [21], Reg-Mod-CS [22], and Weighted- ℓ_1 [25] algorithms are used for comparison, where FISTA- ℓ_1 - ℓ_1 denotes the ℓ_1 - ℓ_1 method that uses only one side information (z_1). The support estimate of x that is considered as prior information by the Mod-CS, Reg-Mod-CS, and Weighted- ℓ_1 algorithms is given by the support of z_1 . Let $m_{3-\ell_1}$, $m_{2-\ell_1}$, $m_{1-\ell_1}$, $m_{\ell_1-\ell_1}$, $m_{\omega\ell_1}$, m_{modCS} , and m_{ℓ_1} denote the bounds of RAMSI-3- ℓ_1 , RAMSI-2- ℓ_1 , RAMSI-1- ℓ_1 , FISTA- ℓ_1 - ℓ_1 , Weighted- ℓ_1 , Mod-CS, and FISTA.

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Fig. 2(a) depicts the Pr(success) versus the number of measurements and Table 1 [Column 2] depicts the bounds for $s_j = 64$. The results show clearly that RAMSI-3- ℓ_1 gives the sharpest bound and the highest recovery performance. For such good quality of side information, the accurate support

Bound	$s_{j} = 64$	$s_j = 256$	$s_j = 352$	
$m_{3-\ell_1}$	208	414	414	
$m_{2-\ell_1}$	229	451	451	
$m_{1-\ell_1}$	272	520	519	
$m_{\ell_1 - \ell_1}$ [30]	259	485	523	
$m_{\omega\ell_1} \equiv m_{ m modCS} \ [21, \ 25]$	253	911	NA	
m_{ℓ_1} [7]	713	713	713	

Table 1: Measurement bounds for $s_0 = 128$ and different $s_j = 64, 256, 352$ values of $s_1 = s_2 = s_3 = s_j$

estimate leads to that the bound $m_{\omega\ell_1} \equiv m_{\text{modCS}}$ is better than $m_{1-\ell_1}$, $m_{\ell_1-\ell_1}$, and m_{ℓ_1} [see Column 2 in Table 1]. Furthermore, the performance of RAMSI-2- ℓ_1 is higher than those of RAMSI-1- ℓ_1 and FISTA- ℓ_1 - ℓ_1 . In this scenario, where the side information is of high quality, FISTA- ℓ_1 - ℓ_1 outperforms our RAMSI-1- ℓ_1 method. Furthermore, Reg-Mod-CS outperforms RAMSI-1- ℓ_1 , FISTA- ℓ_1 - ℓ_1 , Mod-240 CS, and Weighted- ℓ_1 . In addition, the weighted- ℓ_1 method performs better than FISTA but worse than the remaining schemes. Hence, in this case, the use of equal weights—as in FISTA- ℓ_1 - ℓ_1 -leads to a higher performance than using adaptive weights with only one side information signal, as it is the case for RAMSI-1- ℓ_1 . This can be explained by comparing the $m_{1-\ell_1}$ and the $m_{\ell_1-\ell_1}$ bounds: It is clear that $\bar{a}_{1-\ell_1}$ [see (22a)] is greater than \bar{h} [see (9b)]. By combining this observation with the small value of s_1 —due to the good-quality side information signal z_1 —, which in turn results in a small $\bar{s}_{\ell_1-\ell_1}$ value, explains why the bound as well as the recovery performance of FISTA- ℓ_1 - ℓ_1 are better than those of RAMSI-1- ℓ_1 [see magenta and black lines in Fig. 2(a)]. We can conclude that by exploiting multiple side information signals we can obtain the best performance and when dealing with only one good-quality side information signal we may choose to reconstruct using equal weights. 250

Figs. 2(b) and 2(c) present the reconstruction performance versus the number of measurements and Table 1 [Columns 3 and 4] present the bounds when the side information signals are less correlated with the signal of interest, that is, $s_j = 256$ and $s_j = 352$. In this scenario, all RAMSI configurations outperform the FISTA, FISTA- ℓ_1 - ℓ_1 , Weighted- ℓ_1 , Reg-Mod-CS, and Mod-CS methods. The performance of RAMSI-1- ℓ_1 is better than those of FISTA- ℓ_1 - ℓ_1 , Weighted- ℓ_1 , Reg-Mod-CS, and Mod-CS. In Table 1 [Column 3], we also see that the bound $m_{\omega \ell_1} \equiv m_{\text{modCS}}$ in (6) is the worst since the support estimate s_e in (6) is high. In Table 1 [Column 4] the bound for Modified-CS is not available since the support estimate's error $s_e > s_0$ is out of the consideration of the bound $m_{\omega \ell_1}$ in (6). Interestingly, in Fig. 2(c), we observe that the accuracy of FISTA- ℓ_1 - ℓ_1 is worse than that of FISTA, i.e., the side information \mathbf{z}_1 does not help, however, RAMSI-1- ℓ_1 still outperforms FISTA. These results highlight the drawback of the ℓ_1 - ℓ_1 minimization method when the side information is of poor quality. Despite

using poor side information signals, all RAMSI versions achieve better results than FISTA due to the proposed re-weighted $n-\ell_1$ minimization algorithm. We observe that the performance of RAMSI-2- ℓ_1 and RAMSI-3- ℓ_1 is slightly worse than that of RAMSI-1- ℓ_1 . These small penalties are due to the poor quality of the side information signals, which has an impact on the iterative update process.

5.2.2. Side Information Quality-Dependence Analysis

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We now assess the impact of the side information quality on the number of measurements required to successfully reconstruct the target signal. We consider different values for $s_1 = s_2 = s_3$, ranging from 20 to 400 (where s_j is the support size of the difference vector $\boldsymbol{x} - \boldsymbol{z}_j$). The signal \boldsymbol{x} has $s_0 = 128$ and it is generated as in Sec. 5.1. For a fixed value of $s_1 = s_2 = s_3$, we report the number of measurements required by RAMSI-3- ℓ_1 , RAMSI-2- ℓ_1 , RAMSI-1- ℓ_1 , FISTA- ℓ_1 - ℓ_1 , FISTA to achieve a probability of successful recovery bounded as Pr(success) ≥ 0.98 .

Fig. 3 shows that the recovery performance is clearly improved when leveraging side information; and the more signals are considered the better the performance. However, when the support size of $x - z_j$ meets that of the target signal, the side information does not improve the performance 275 anymore. When $s_j > s_0$, the performance and the measurement bound of FISTA- ℓ_1 - ℓ_1 are increasing with the value of s_1 . Specifically, when $s_1 > 315$, the performance of FISTA- ℓ_1 - ℓ_1 is worse than that of FISTA, thereby illustrating the limitations of ℓ_1 - ℓ_1 minimization when the side information is of poor quality. As shown in Fig. 3, the performance of RAMSI-both in terms of the bounds and the practical results—is robust against poor-quality side information. The bounds are sharper 280 when the number of side information signals increases and remain approximately constant when s_i increases after a threshold (indicating increasing-poor side information quality). When $s_j > 300$, the number of measurements of RAMSI-3- ℓ_1 and RAMSI-2- ℓ_1 are slightly worse than those of RAMSI-1- ℓ_1 , approaching the number of measurements of FISTA [this behavior was also observed in Fig. 2(c)]. To address this issue, we can adaptively select the best among the RAMSI configurations. For 285 instance, when one side information is available, we can choose to use equal weights in case of good side information. Furthermore, we can ensure that RAMSI's performance is not worse than FISTA by weighing dominantly on the signal rather than on the side information signals. To do so, we need to make an estimate of the quality of the side information signals, which is left as a topic for future research. 290

6. Conclusion

We proposed the RAMSI algorithm and established measurement bounds for the problem of signal recovery under a weighted $n-\ell_1$ minimization framework. RAMSI incorporates multiple side information signals in the problem of sparse signal recovery and iteratively weights them so as to optimize the



Figure 3: Number of measurements vs. number of nonzeros $s_1 = s_2 = s_3$ of side information signals given the source $s_0 = 128$ for RAMSI using one, two, three side information signals.

reconstruction performance. The bounds confirm the advantage of RAMSI in utilizing multiple side information signals to significantly reduce the number of measurements and to deal with variations in the quality of the side information. We experimentally assessed the bounds and the performance of RAMSI against state-of-the-art methods using numerical sparse signals. The results showed that our bounds are sharper than existing bounds and that RAMSI outperforms the conventional CS method, the recent ℓ_1 - ℓ_1 minimization method as well as the Modified-CS, regularized Modified-CS, and weighted ℓ_1 300 minimization methods.

Appendix A. Background on Measurement Bounds

We summarise some key definitions and conditions in convex optimization and linear inverse problems [11, 12], which are used in the derivation of the measurement bounds for the proposed weighted $n-\ell_1$ minimization approach.

Appendix A.0.1. Convex Cone

A convex cone $C \subset \mathbb{R}^n$ is a convex set that satisfies $C = \tau C$, $\forall \tau \geq 0$ [12]. For the cone $C \subset \mathbb{R}^n$, a *polar cone* C° is the set of outward normals of C, defined by

$$C^{\circ} := \{ \boldsymbol{u} \in \mathbb{R}^n : \boldsymbol{u}^{\mathrm{T}} \boldsymbol{x} \le 0, \ \forall \boldsymbol{x} \in C \}.$$
(A.1)

A descent cone [12, Definition 2.7] $\mathcal{D}(g, \boldsymbol{x})$, alias tangent cone [11], of a convex function $g := \mathbb{R}^n \to \mathbb{R}$ at a point $x \in \mathbb{R}^n$ —at which g is not increasing—is defined as

$$\mathcal{D}(g, \boldsymbol{x}) := \bigcup_{\tau \ge 0} \{ \boldsymbol{y} \in \mathbb{R}^n : g(\boldsymbol{x} + \tau \boldsymbol{y}) \le g(\boldsymbol{x}) \},$$
(A.2)

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where [] denotes the union operator.

Appendix A.0.2. Gaussian Width

The Gaussian width [11] is a summary parameter for convex cones; it is used to measure the aperture of a convex cone. For a convex cone $C \subset \mathbb{R}^n$, considering a subset $C \cap \mathbb{S}^{n-1}$ where $\mathbb{S}^{n-1} \subset \mathbb{R}^n$ is a unit sphere, the *Gaussian width* [11, Definition 3.1] is defined as

$$\omega(C) := \mathbb{E}_{\mathbf{g}}[\sup_{\mathbf{u} \in C \cap \mathbb{S}^{n-1}} \mathbf{g}^{\mathrm{T}} \mathbf{u}].$$
(A.3)

where $\mathbf{g} \sim \mathcal{N}(0, \mathbf{I}_n)$ is a vector of *n* independent, zero-mean, and unit-variance Gaussian random variables and $\mathbb{E}_{\mathbf{g}}[\cdot]$ denotes the expectation with respect to \mathbf{g} . The Gaussian width [11, Proposition 3.6] can further be bounded as

$$\omega(C) \le \mathbb{E}_{\boldsymbol{g}} [\operatorname{dist}(\boldsymbol{g}, C^{\circ})], \tag{A.4}$$

where $dist(\mathbf{g}, C^{\circ})$ denotes the Euclidean distance of \mathbf{g} with respect to the set C° , which is in turn defined as

$$\operatorname{dist}(\boldsymbol{g}, C^{\circ}) := \min_{\boldsymbol{u}} \{ \|\boldsymbol{g} - \boldsymbol{u}\|_2 : \boldsymbol{u} \in C^{\circ} \}.$$
(A.5)

Recently, a summary parameter called the *statistical dimension* $\delta(C)$ of the cone C [12] is used to estimate the convex cone [12, Theorem 4.3]. The statistical dimension is expressed in terms of the polar cone C° as [12, Proposition 3.1]

$$\delta(C) := \mathbb{E}_{\mathbf{g}}[\operatorname{dist}^2(\mathbf{g}, C^\circ)].$$
(A.6)

From (A.4) and (A.6), we can derive the following inequality

$$\omega^2(C) \le \delta(C),\tag{A.7}$$

which gives a convenient bound for the Gaussian width that will be used in our computations.

310 Appendix A.O.3. Measurement Condition

An optimality condition [11, Proposition 2.1], [12, Fact 2.8] for linear inverse problems states that x_0 is the unique solution of (2) if and only if

$$\mathcal{D}(g, \boldsymbol{x}_0) \cap \operatorname{null}(\boldsymbol{\Phi}) = \{\boldsymbol{0}\},\tag{A.8}$$

where null(Φ):= { $x \in \mathbb{R}^n : \Phi x = 0$ } is the null space of Φ . We consider the number of measurements m required to successfully recover a given signal $x_0 \in \mathbb{R}^n$. Corollary 3.3 in [11] states that, given a measurement vector $y = \Phi x_0, x_0$ is the unique solution of (2) with probability at least 1-exp $(-\frac{1}{2}(\sqrt{m}-\omega(\mathcal{D}(g,x_0)))^2)$ provided that $m \ge \omega^2(\mathcal{D}(g,x_0)) + 1$. Using (A.7), we can write this condition as

$$m \ge \delta(\mathcal{D}(g, \boldsymbol{x}_0)) + 1. \tag{A.9}$$

Appendix A.0.4. Bound on the Measurement Condition

Using (A.6), the statistical dimension $\delta(\mathcal{D}(g, \boldsymbol{x}))$ of a descent cone $\mathcal{D}(g, \boldsymbol{x})$ is

$$\delta(\mathcal{D}(g, \boldsymbol{x})) = \mathbb{E}_{\boldsymbol{g}} \left[\operatorname{dist}^{2}(\boldsymbol{g}, \mathcal{D}(g, \boldsymbol{x})^{\circ}) \right], \tag{A.10}$$

where $\mathcal{D}(g, \boldsymbol{x})^{\circ}$ is the polar cone of $\mathcal{D}(g, \boldsymbol{x})$ defined in (A.1). Let us consider that the subdifferential ∂g [56] of a convex function g at a point $\boldsymbol{x} \in \mathbb{R}^n$ is given by $\partial g(\boldsymbol{x}) := \{\boldsymbol{u} \in \mathbb{R}^n : g(\boldsymbol{y}) \ge g(\boldsymbol{x}) + \boldsymbol{u}^{\mathrm{T}}(\boldsymbol{y} - \boldsymbol{x})\}$ for all $\boldsymbol{y} \in \mathbb{R}^n\}$. From (A.10) and [12, Proposition 4.1], we obtain a lower bound on $\delta(\mathcal{D}(g, \boldsymbol{x}))$ as

$$\delta(\mathcal{D}(g, \boldsymbol{x})) = \mathbb{E}_{\boldsymbol{g}} \Big[\min_{\tau \ge 0} \operatorname{dist}^2(\boldsymbol{g}, \tau \cdot \partial g(\boldsymbol{x})) \Big] \le \min_{\tau \ge 0} \mathbb{E}_{\boldsymbol{g}} \Big[\operatorname{dist}^2(\boldsymbol{g}, \tau \cdot \partial g(\boldsymbol{x})) \Big].$$
(A.11)

In short, we conclude the following proposition.

Proposition Appendix A.1 (Measurement bound for a convex norm function). In order to obtain the measurement bound for the recovery condition, $m \ge U_g + 1$, we calculate the quantity U_g given a convex norm function $g := \mathbb{R}^n \to \mathbb{R}$ by

$$U_g = \min_{\tau \ge 0} \mathbb{E}_{\boldsymbol{g}} \Big[\operatorname{dist}^2(\boldsymbol{g}; \tau \cdot \partial g(\boldsymbol{x})) \Big].$$
(A.12)

Appendix B. Proof of Proposition 3.1

Replacing $g(\mathbf{x}) = \lambda \sum_{j=0}^{J} \|\mathbf{W}_j(\mathbf{x} - \mathbf{z}_j)\|_1$ in the definition of the proximal operator in (4), we obtain

$$\Gamma_{\frac{1}{L}g}(\boldsymbol{x}) = \arg\min_{\boldsymbol{v}\in\mathbb{R}^n} \left\{ \frac{\lambda}{L} \sum_{j=0}^{J} \|\mathbf{W}_j(\boldsymbol{v}-\boldsymbol{z}_j)\|_1 + \frac{1}{2} \|\boldsymbol{v}-\boldsymbol{x}\|_2^2 \right\}.$$
(B.1)

Both terms in (B.1) are separable in v and thus, we can minimize each element v_i of v individually as

$$\Gamma_{\overline{L}g}(x_i) = \operatorname*{arg\,min}_{v_i \in \mathbb{R}} \Big\{ h(v_i) = \frac{\lambda}{L} \sum_{j=0}^J w_{ji} |v_i - z_{ji}| + \frac{1}{2} (v_i - x_i)^2 \Big\}.$$
(B.2)

To solve (B.2) we need to derive $\frac{\partial h(v_i)}{\partial v_i}$. Recall from Definition 4.3 that, without loss of generality, we assume that $-\infty = z_{(-1)i} \leq z_{0i} \leq z_{1i} \leq \cdots \leq z_{Ji} \leq z_{(J+1)i} = \infty$. When $v_i \in (z_{ri}, z_{(r+1)i})$ with $r \in \{-1, \ldots, J\}, h(v_i)$ is differentiable and $\frac{\partial h(v_i)}{\partial v_i}$ is calculated as

$$\frac{\partial h(v_i)}{\partial v_i} = \frac{\lambda}{L} \sum_{j=0}^{J} w_{ji} \operatorname{sign}(v_i - z_{ji}) + (v_i - x_i),$$
(B.3)

where sign(.) is the sign function. Using the definition of the $\mathfrak{b}(.)$ function in (17), we have that $\operatorname{sign}(v_i - z_{ji}) = (-1)^{\mathfrak{b}(r < j)}$ and thus, we can rewrite (B.3) as

$$\frac{\partial h(v_i)}{\partial v_i} = \frac{\lambda}{L} \sum_{j=0}^J w_{ji} (-1)^{\mathfrak{b}(r(B.4)$$

By setting $\frac{\partial h(v_i)}{\partial v_i} = 0$, we obtain:

$$v_i = x_i - \frac{\lambda}{L} \sum_{j=0}^{J} w_{ji} (-1)^{\mathfrak{b}(r < j)}.$$
 (B.5)

Since $v_i \in (z_{ri}, z_{(r+1)i})$, via (B.5) we have that

$$z_{ri} + \frac{\lambda}{L} \sum_{j=0}^{J} w_{ji} (-1)^{\mathfrak{b}(r(B.6)$$

In the following Lemma Appendix B.1, we prove that, in the remaining range value of x_i , namely, in the case that

$$z_{ri} + \frac{\lambda}{L} \sum_{j=0}^{J} w_{ji} (-1)^{\mathfrak{b}(r-1$$

the minimum of $h(v_i)$ is obtained when $v_i = z_{ri}$.

Lemma Appendix B.1. Given that x_i is bounded as in (B.7), the function $h(v_i)$ defined in (B.2) is minimized when $v_i = z_{ri}$.

Proof. Let us express $h(v_i)$ as

$$h(v_i) = \frac{\lambda}{L} \sum_{j=0}^{J} w_{ji} |(v_i - z_{ri}) - (z_{ji} - z_{ri})| + \frac{1}{2} [(v_i - z_{ri}) - (x_i - z_{ri})]^2.$$
(B.8)

Applying the inequality $|a - b| \ge |a| - |b|$, with $a, b \in \mathbb{R}$, on the first summation term and expanding the second term, we obtain:

$$h(v_i) \ge \frac{\lambda}{L} \sum_{j=0}^{J} w_{ji} |v_i - z_{ri}| - \frac{\lambda}{L} \sum_{j=0}^{J} w_{ji} |z_{ji} - z_{ri}| + \frac{1}{2} (v_i - z_{ri})^2 - (v_i - z_{ri}) \cdot (x_i - z_{ri}) + \frac{1}{2} (x_i - z_{ri})^2.$$
(B.9)

Using the basic inequality $-(v_i - z_{ri}) \cdot (x_i - z_{ri}) \ge -|v_i - z_{ri}| \cdot |x_i - z_{ri}|$, we can express (B.9) as

$$h(v_i) \ge |v_i - z_{ri}| \frac{\lambda}{L} \sum_{j=0}^{J} w_{ji} - |v_i - z_{ri}| \cdot |x_i - z_{ri}| + \frac{1}{2} (v_i - z_{ri})^2 - \frac{\lambda}{L} \sum_{j=0}^{J} w_{ji} |z_{ji} - z_{ri}| + \frac{1}{2} (x_i - z_{ri})^2.$$
(B.10)

At this point, via (B.7) we obtain

$$-\frac{\lambda}{L}\sum_{j=0}^{J}w_{ji} \le x_i - z_{ri} \le \frac{\lambda}{L}\sum_{j=0}^{J}w_{ji} \Rightarrow |x_i - z_{ri}| \le \frac{\lambda}{L}\sum_{j=0}^{J}w_{ji}.$$
 (B.11)

We now observe that the part including v_i in the right hand side of (B.10) can be written as

$$|v_i - z_{ri}| \left(\frac{\lambda}{L} \sum_{j=0}^{J} w_{ji} - |x_i - z_{ri}|\right) + \frac{1}{2} (v_i - z_{ri})^2.$$
(B.12)

Taking into account (B.11), the expression in (B.12) and, in turn $h(v_i)$, are minimized when $v_i = z_{ri}$.

In summary, via (B.5) and Lemma Appendix B.1, we obtain

$$\Gamma_{\frac{1}{L}g}(x_i) = \begin{cases} x_i - \frac{\lambda}{L} \sum_{j=0}^{J} w_{ji}(-1)^{\mathfrak{b}(r(B.13)$$

³²⁰ Appendix C. Proof of Theorem 4.4 and Corollary 4.4.1

We begin by stating some important results that help us proving Theorem 4.4 and Corollary 4.4.1. Recall that the probability density of the normal distribution $\mathcal{N}(0,1)$ with zero-mean and unit variance $\psi(x)$ is given by

$$\psi(x) := \frac{1}{\sqrt{2\pi}} e^{-x^2/2}.$$
(C.1)

Our formulations consider the following inequality, which is stated in [30], that is

$$\frac{(1-x^{-1})}{\sqrt{\pi\log(x)}} \le \frac{1}{\sqrt{2\pi}} \le \frac{2}{5},$$
(C.2)

for all x > 1. Moreover, adhering to the formulations in [30], we use the following expressions in our derivations:

$$\mathcal{A}(x) := \frac{1}{\sqrt{2\pi}} \int_{x}^{\infty} (v - x)^2 e^{-v^2/2} \mathrm{d}v,$$
(C.3a)

$$\mathcal{B}(x) := \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{x} (v - x)^2 e^{-v^2/2} \mathrm{d}v, \qquad (C.3b)$$

for which we have that

$$\mathcal{A}(0) = \mathcal{B}(0) = 1/2.$$
 (C.4)

When $x \neq 0$, the following inequalities have been derived in [30]:

$$\mathcal{A}(x) \le \begin{cases} \psi(x)/x, & x > 0\\ x^2 + 1, & x < 0 \end{cases}$$
(C.5a)

$$\mathcal{B}(x) \le \begin{cases} -\psi(x)/x, & x < 0\\ x^2 + 1, & x > 0. \end{cases}$$
(C.5b)

Lemma Appendix C.1. Given $x \in (0, 1]$ and $\tau > 0$ for $\psi(x)$ given in (C.1), we have

$$\frac{\psi(\tau x)}{\tau x} \le \frac{1}{\sqrt{2\pi}} \frac{1 - x^2}{\tau x} + x \frac{\psi(\tau)}{\tau}.$$
(C.6)

Proof. Using (C.1) the left hand side of (C.6) becomes

$$\frac{\psi(\tau x)}{\tau x} = \frac{1}{\sqrt{2\pi}} \frac{e^{-\tau^2 x^2/2}}{\tau x}.$$
(C.7)

Applying Bernoulli's inequality on $e^{-\tau^2 x^2/2}$ leads to

$$e^{-\tau^2 x^2/2} = \left(1 + (e^{-\tau^2/2} - 1)\right)^{x^2} \le 1 + x^2(e^{-\tau^2/2} - 1),$$
 (C.8)

where $0 < x \le 1$ and $(e^{-\tau^2/2}-1) > -1$ given that $\tau > 0$. Combining (C.7) and (C.8) leads to the proof. \Box

Proof of Theorem 4.4. We derive the bound based on Proposition Appendix A.1 by firstly computing the subdifferential $\partial g(\mathbf{x})$ and then the distance between the standard normal vector \mathbf{g} to $\partial g(\mathbf{x})$. Under Definitions 4.2 and 4.3, let us recall that there are d_i consecutive zero elements at any $i \in \{p+1,\ldots,q\}$, namely, $\{x_i - z_{ji}\}_{j=l_i}^{l_i+d_i-1} = 0$, with $l_i \in \{0,\ldots,J\}$ being an auxiliary index indicating the start of the zero positions. Under Definitions 4.1, 4.2, 4.3, and the weights in (14), the elements of the vectors \mathbf{u} in the subdifferential $\partial g(\mathbf{x})$ are expressed as

$$u_{i} = a_{i}, \qquad i \in \{1, \dots, p\}, \\ u_{i} \in [b_{i} - c_{i}, b_{i} + c_{i}], \qquad i \in \{p + 1, \dots, q\}, \\ u_{i} \in \left[-\sum_{j=0}^{J} w_{ji}, \sum_{j=0}^{J} w_{ji}\right] = [-1, 1], \quad i \in \{q + 1, \dots, n\}.$$
(C.9)

where

$$a_i = \sum_{j=0}^{J} w_{ji} (-1)^{\mathfrak{b}(r_i < j)}, \tag{C.10a}$$

$$b_i = \sum_{j \notin \{l_i, \dots, l_i + d_i - 1\}} w_{ji} (-1)^{\mathfrak{b}(r_i < j)}, \tag{C.10b}$$

$$c_i = \sum_{j=l_i}^{l_i+d_i-1} w_{ji} = d_i \frac{\eta_i}{\epsilon}$$
(C.10c)

Using (A.5) and (C.9), we can compute the distance from the standard normal vector \mathbf{g} to the subdifferential $\partial g(\mathbf{x})$ as

$$\operatorname{dist}^{2}(\boldsymbol{g}, \tau \cdot \partial g(\boldsymbol{x})) = \sum_{i=1}^{p} (g_{i} - \tau a_{i})^{2} + \sum_{i=p+1}^{q} \left(\mathcal{P}^{2}(g_{i} - \tau(b_{i} + c_{i})) + \mathcal{P}^{2}(-g_{i} + \tau(b_{i} - c_{i})) \right) + \sum_{i=q+1}^{n} \mathcal{P}^{2}(|g_{i}| - \tau),$$
(C.11)

where $\mathcal{P}(a) := \max\{a, 0\}$ returns the maximum value between $a \in \mathbb{R}$ and 0. Taking the expectation of (C.11) with respect to g delivers

$$\mathbb{E}_{\mathbf{g}}[\operatorname{dist}^{2}(\mathbf{g},\tau\cdot\partial g(\mathbf{x}))] = p + \tau^{2}\sum_{i=1}^{p}a_{i}^{2} + \frac{1}{\sqrt{2\pi}}\sum_{i=p+1}^{q}\int_{\tau(b_{i}+c_{i})}^{\infty}(v-\tau(b_{i}+c_{i}))^{2}\mathrm{e}^{-v^{2}/2}\mathrm{d}v + \frac{1}{\sqrt{2\pi}}\sum_{i=p+1}^{q}\int_{-\infty}^{\tau(b_{i}-c_{i})}(v-\tau(b_{i}-c_{i}))^{2}\mathrm{e}^{-v^{2}/2}\mathrm{d}v + \sqrt{\frac{2}{\pi}}\sum_{i=q+1}^{n}\int_{\tau}^{\infty}(v-\tau)^{2}\mathrm{e}^{-v^{2}/2}\mathrm{d}v. \quad (C.12)$$

Replacing the expressions in (C.3a), (C.3b) in (C.12) gives

$$\mathbb{E}_{\mathbf{g}}[\operatorname{dist}^{2}(\mathbf{g},\tau\cdot\partial g(\mathbf{x}))] = p + \tau^{2}\sum_{i=1}^{p}a_{i}^{2} + \sum_{i=p+1}^{q}\mathcal{A}(\tau(b_{i}+c_{i})) + \sum_{i=p+1}^{q}\mathcal{B}(\tau(b_{i}-c_{i})) + 2\sum_{i=q+1}^{n}\mathcal{A}(\tau). \quad (C.13)$$

At this point, it is worth emphasizing the advantage of the adaptive weights [see (14)] in the proposed method, the values of which depend on the correlation of the side information with the target signal. Focusing on a given index $i \in \{p + 1, \ldots, q\}$, let us observe the weight contribution in the expressions of b_i and c_i , defined in (C.10b) and (C.10c), respectively. The weights w_{ji} in c_i are considerably higher than those in b_i , as in the former case the side information signals z_{ji} are equal to the source x_i . Moreover, we recall that a small positive parameter ϵ is introduced in the denominator of the weights so as to avoid division by zero when $x_i = z_{ji}$. The ϵ parameter can ensure that c_i in (C.10c) is always greater than $|b_i|$ in (C.10b). Hence, we always have $b_i + c_i > 0$ and $b_i - c_i < 0$. With these observations we conclude that the arguments of $\mathcal{A}(\cdot)$ and $\mathcal{B}(\cdot)$ in (C.13) are respectively positive and negative. Applying inequality (C.5a) for x > 0 on the expression of $\mathcal{A}(\cdot)$ as well as inequality (C.5b) for x < 0 on the expression of $\mathcal{B}(\cdot)$, we obtain the following bound for the U_g quantity [which is defined in (A.12)]:

$$U_{g_{n-\ell_1}} \le \min_{\tau \ge 0} \Big\{ p + \tau^2 \sum_{i=1}^p a_i^2 + \sum_{i=p+1}^q \Big(\frac{\psi(\tau(b_i + c_i))}{\tau(b_i + c_i)} + \frac{\psi(\tau(c_i - b_i))}{\tau(c_i - b_i)} \Big) + 2 \sum_{i=q+1}^n \frac{\psi(\tau)}{\tau} \Big\},$$
(C.14)

where $\psi(\cdot)$ is the zero-mean, unit-variance normal distribution defined in (C.1). Applying (C.6) on the second sum in (C.14) gives

$$U_{g_{n-\ell_1}} \le \min_{\tau \ge 0} \Big\{ p + \tau^2 \sum_{i=1}^p a_i^2 + \sum_{i=p+1}^q \Big(\frac{1 - (b_i + c_i)^2}{\tau \sqrt{2\pi} (b_i + c_i)} + \frac{1 - (c_i - b_i)^2}{\tau \sqrt{2\pi} (c_i - b_i)} + c_i \frac{2\psi(\tau)}{\tau} \Big) + 2(n-q) \frac{\psi(\tau)}{\tau} \Big\}.$$
(C.15)

$$\Rightarrow U_{g_{n-\ell_1}} \le \min_{\tau \ge 0} \Big\{ p + \tau^2 \sum_{i=1}^p a_i^2 + \sum_{i=p+1}^q \frac{1}{\sqrt{2\pi}} \frac{2c_i}{\tau} \Big(\frac{1}{c_i^2 - b_i^2} - 1 \Big) + \Big((n-q) + \sum_{i=p+1}^q c_i \Big) 2 \frac{\psi(\tau)}{\tau} \Big\}.$$
(C.16)

From the definitions of b_i and c_i in (C.10b) and (C.10c), respectively, we have

 $\sum_{j \notin \{l_i, \dots, l_i + d_i - 1\}} w_{ji} = 1 - c_i,$ (C.17)

where we used the constraint $\sum_{j=0}^{J} w_{ji} = 1$. Hence, the second summation term in (C.16) is bounded as

$$\sum_{i=p+1}^{q} \frac{1}{\sqrt{2\pi}} \frac{2c_i}{\tau} \left(\frac{1}{c_i^2 - b_i^2} - 1 \right) \le \sum_{i=p+1}^{q} \frac{4}{\sqrt{2\pi}\tau} \frac{c_i}{2c_i - 1} (1 - c_i)$$
(C.18)

$$\leq \frac{4 \cdot \min\{c_i\}}{\sqrt{2\pi\tau}(2 \cdot \min\{c_i\} - 1)} \sum_{i=p+1}^{q} (1 - c_i),$$
(C.19)

where (C.18) is obtained by using that $b_i \leq 1 - c_i$ and (C.19) holds from the fact that the term $\frac{c_i}{2c_i-1}$ is maximized when c_i is minimized (recall that $c_i > 0$). For simplicity, let us denote

$$\bar{a}_{n-\ell_1} = \sum_{i=1}^p a_i^2, \tag{C.20a}$$

$$\kappa_{n-\ell_1} = \frac{4 \cdot \min\{c_i\}}{\sqrt{2\pi\tau}(2 \cdot \min\{c_i\} - 1)},\tag{C.20b}$$

$$\bar{s}_{n-\ell_1} = q - \sum_{i=p+1}^{q} c_i = p + \sum_{i=p+1}^{q} (1 - c_i).$$
 (C.20c)

Substituting the quantities of (C.20a), (C.20b), (C.20c), and using inequality (C.19) in (C.16) gives

$$U_{g_{n-\ell_1}} \le \min_{\tau \ge 0} \left\{ p + \bar{a}_{n-\ell_1} \tau^2 + \kappa_{n-\ell_1} (\bar{s}_{n-\ell_1} - p) + (n - \bar{s}_{n-\ell_1}) \frac{2\psi(\tau)}{\tau} \right\}$$

which using the definition of $\psi(\cdot)$ can be written as

$$U_{g_{n-\ell_1}} \le \min_{\tau \ge 0} \left\{ \bar{a}_{n-\ell_1} \tau^2 + (n - \bar{s}_{n-\ell_1}) \frac{2}{\sqrt{2\pi}} \frac{e^{-\frac{\tau^2}{2}}}{\tau} + \bar{s}_{n-\ell_1} + (\kappa_{n-\ell_1} - 1)(\bar{s}_{n-\ell_1} - p) \right\}.$$
 (C.21)

To derive a bound as a function of the source signal x, we need to select a parameter $\tau > 0$. Setting $\tau = \sqrt{2 \log(n/\bar{s}_{n-\ell_1})}$ yields

$$U_{g_{n-\ell_1}} \le 2\bar{a}_{n-\ell_1} \log \frac{n}{\bar{s}_{n-\ell_1}} + \frac{\bar{s}_{n-\ell_1}(1-\bar{s}_{n-\ell_1}/n)}{\sqrt{\pi \log(n/\bar{s}_{n-\ell_1})}} + \bar{s}_{n-\ell_1} + \delta_{n-\ell_1}, \tag{C.22}$$

where

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$$\delta_{n-\ell_1} = (\kappa_{n-\ell_1} - 1)(\bar{s}_{n-\ell_1} - p), \tag{C.23}$$

and where we have replaced the selected value of τ in (C.20b), thereby obtaining the $\kappa_{n-\ell_1}$ definition reported in (23).

Applying inequality (C.2) on the second term of the right hand side of (C.22) gives

$$U_{g_{n-\ell_1}} \le 2\bar{a}_{n-\ell_1} \log_1 \frac{n}{\bar{s}_{n-\ell_1}} + \frac{7}{5} \bar{s}_{n-\ell_1} + \delta_{n-\ell_1}.$$
(C.24)

Bearing in mind that $m_{n-\ell_1} \ge U_{g_{n-\ell_1}} + 1$ and by combining (C.24) with (C.20), (C.10), and (C.23) leads to the proof.

Proof of Corollary 4.4.1. We start with Relation (a). Under the conditions that $\mathbf{W}_0 = \mathbf{I}_n$ and $\mathbf{W}_j = \mathbf{0}$ for $j \in \{1, \dots, J\}$, we have that $\bar{s}_{n-\ell_1} = p = s_0$ and $\bar{a}_{n-\ell_1} = p = s_0$, where we used the definitions in (C.20a), (C.20c). Consequently, from (C.23), $\delta_{n-\ell_1} = 0$. Replacing these values of $\bar{a}_{n-\ell_1}$, $\bar{s}_{n-\ell_1}$, $\delta_{n-\ell_1}$ in our bound, defined in (21), leads to the ℓ_1 minimization bound in (5).

To reach Relation (b), given that $\mathbf{W}_0 = \mathbf{W}_1 = \frac{1}{2}\mathbf{I}_n$ and $\mathbf{W}_j = \mathbf{0}$ for $j \in \{2, \ldots, J\}$, let us first denote two subsets, \mathbb{I}_1 and \mathbb{I}_2 , as

$$\mathbb{I}_1 := \Big\{ i \in \{ p+1, \dots, q \} : b_i + c_i = 1, b_i - c_i = 0 \Big\},$$
(C.25a)

$$\mathbb{I}_2 := \Big\{ i \in \{p+1, \dots, q\} : b_i + c_i = 0, b_i - c_i = -1 \Big\}.$$
 (C.25b)

Via the definitions of b_i and c_i —see (C.10b) and (C.10c), respectively—and since $i \in \{p + 1, \ldots, q\}$, we observe that $i \in \mathbb{I}_1$ or $i \in \mathbb{I}_2$.

Replacing (C.25a) and (C.25b) in (C.13) leads to

$$\mathbb{E}_{\mathbf{g}}[\operatorname{dist}^{2}(\mathbf{g},\tau\cdot\partial g(\mathbf{x}))] = p + \tau^{2}\sum_{i=1}^{p}a_{i}^{2} + \sum_{i\in\mathbb{I}_{1}}\mathcal{A}(\tau) + \sum_{i\in\mathbb{I}_{1}}\mathcal{B}(0) + \sum_{i\in\mathbb{I}_{2}}\mathcal{A}(0) + \sum_{i\in\mathbb{I}_{2}}\mathcal{B}(-\tau) + 2\sum_{i=q+1}^{n}\mathcal{A}(\tau). \quad (C.26)$$

By combining (C.4), (C.5a), and (C.5b) with (C.26), we obtain the following bound for the U_g quantity [defined in (A.12)]:

$$U_{g_{\ell_1-\ell_1}} \le \min_{\tau \ge 0} \Big\{ p + \tau^2 \sum_{i=1}^p a_i^2 + \sum_{i=p+1}^q \frac{1}{2} + \sum_{i=p+1}^q \frac{\psi(\tau)}{\tau} + 2\sum_{i=q+1}^n \frac{\psi(\tau)}{\tau} \Big\},\tag{C.27}$$

which can further be elaborated to

$$U_{g_{\ell_1-\ell_1}} \le \min_{\tau \ge 0} \left\{ \tau^2 \sum_{i=1}^p a_i^2 + \frac{1}{2} (p+q) + (2n - (p+q)) \frac{e^{\frac{-\tau^2}{2}}}{\sqrt{2\pi\tau}} \right\}.$$
(C.28)

In the ℓ_1 - ℓ_1 minimization case, there is a single side information signal and no weights, that is, $d_i = 1$ and J = 1 in (20) under Definitions 4.1 and 4.2; thus, we have $p + q = s_0 + s_1$. Combining this result with (9b) gives $\sum_{i=1}^{p} a_i^2 = \bar{h}$.

Let us now denote $\bar{s}_{\ell_1-\ell_1} = \frac{s_0+s_1}{2}$ and set $\tau = \sqrt{2\log(n/\bar{s}_{\ell_1-\ell_1})}$; thus, we have

$$U_{g_{\ell_1-\ell_1}} \le 2\bar{h}\log\frac{n}{\bar{s}_{\ell_1-\ell_1}} + \bar{s}_{\ell_1-\ell_1} + \frac{\bar{s}_{\ell_1-\ell_1}(1 - \frac{\bar{s}_{\ell_1-\ell_1}}{n})}{\sqrt{2\pi\log(\frac{n}{\bar{s}_{\ell_1-\ell_1}})}}.$$
(C.29)

Applying (C.2) on the third term of the right hand side of (C.29) gives

$$U_{g_{\ell_1-\ell_1}} \le 2\bar{h}\log\frac{n}{\bar{s}_{\ell_1-\ell_1}} + \frac{7}{5}\bar{s}_{\ell_1-\ell_1}.$$
(C.30)

Finally, we obtain the ℓ_1 - ℓ_1 minimization bound [29, 30, 34] in (8) as

$$m_{\ell_1-\ell_1} \ge 2\bar{h}\log\frac{n}{\bar{s}_{\ell_1-\ell_1}} + \frac{7}{5}\bar{s}_{\ell_1-\ell_1} + 1,$$
 (C.31)

where $\bar{s}_{\ell_1-\ell_1} = \frac{s_0+s_1}{2} = s_0 + \frac{\xi}{2}$, with ξ defined in (9a).

Appendix D. Supporting Lemmas

Lemma Appendix D.1. The value of $\delta_{n-\ell_1}$ in the definition of the bound for weighted $n-\ell_1$ minimization—given in (21)—is negative.

Proof. Recall the definition of $\delta_{n-\ell_1}$ given in (C.23). By the definition of $\bar{s}_{n-\ell_1}$ in (C.20c), it is clear that $(\bar{s}_{n-\ell_1}-p) > 0$; this is because $c_i < 1$ [see (C.10c)]. Hence, the sign of $\delta_{n-\ell_1}$ depends on the term $(\kappa_{n-\ell_1}-1)$. From (C.20b), it is clear that $\kappa_{n-\ell_1}$ depends on c_i , which is defined as

$$c_{i} = d_{i} \left(d_{i} + \sum_{j \notin \{l_{i}, \dots, l_{i} + d_{i} - 1\}} \frac{\epsilon}{|x_{i} - z_{ji}| + \epsilon} \right)^{-1}.$$
 (D.1)

As $\epsilon > 0$ and very small, we can say that $c_i \to 1^-$. As a result, $\kappa_{n-\ell_1}$ is approximately given by

$$\kappa_{n-\ell_1} \approx \frac{4}{\sqrt{2\pi\tau}} \approx \frac{2}{\sqrt{\pi \log(n/\bar{s}_{n-\ell_1})}},\tag{D.2}$$

- where in the proof of Theorem 4.4 we have set that $\tau = \sqrt{2 \log(n/\bar{s}_{n-\ell_1})}$. We observe that $\kappa_{n-\ell_1} < 1$ if $\frac{\bar{s}_{n-\ell_1}}{n} < 0.28$, where $\bar{s}_{n-\ell_1} \approx p \leq \min\{s_j\}$ [see (C.20c) and Definition 4.1]. In Lemma Appendix D.2, we prove that when $\frac{p}{n} < 0.23$ then $\hat{m}_{n-\ell_1}$ in (27) is less than the source dimension, n. Otherwise, the required number of measurements $\hat{m}_{n-\ell_1}$ is higher than n, signifying the failure of the algorithm². Hence, we get that $(\kappa_{n-\ell_1} - 1) < 0$, thereby proving that $\delta_{n-\ell_1} < 0$.
- Lemma Appendix D.2. Given a sparse signal $\boldsymbol{x} \in \mathbb{R}^n$ with $\|\boldsymbol{x}\|_0 = s_0$ and side information signals $\boldsymbol{z}_j \in \mathbb{R}^n$ with $\|\boldsymbol{x} \boldsymbol{z}_j\|_0 = s_j, \forall j \in \{1, \dots, J\}$, the number of measurements required for weighted $n \ell_1$ minimization to recover \boldsymbol{x} —given in (27)—is less than the source dimension n when $\frac{p}{n} < 0.23$.

Proof. Let us assume that the number of measurements in the bound of (27) satisfies the condition $\widehat{m}_{n-\ell_1} < n$, namely, we have

$$2p\log\frac{n}{p} + \frac{7}{5}p < n. \tag{D.3}$$

By substituting $\gamma = \frac{n}{p}, \gamma \in \mathbb{R}$, in (D.3) we get

$$2\log\gamma + \frac{7}{5} < \gamma. \tag{D.4}$$

By setting $f(\gamma) = \gamma - 2\log\gamma - \frac{7}{5}$ and using calculus we can show that $f(\gamma) > 0$ when $\gamma > 4.33$, or else, $\frac{p}{n} < 0.23$.

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