



# Do green technology innovations contribute to carbon dioxide emission reduction? Empirical evidence from patent data

Kerui Du<sup>a,\*</sup>, Pengzhen Li<sup>b</sup>, Zheming Yan<sup>c,\*</sup>

<sup>a</sup> School of Management, China Institute for Studies in Energy Policy, Collaborative Innovation Center for Energy Economics and Energy Policy, Xiamen University, Xiamen, Fujian 361005, PR China

<sup>b</sup> Center for Economic Research, Shandong University, Jinan 250100, China

<sup>c</sup> International Business School, Shaanxi Normal University, Xi'an 710119, China



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## ABSTRACT

This paper investigates the impact of green technology innovations on carbon dioxide (CO<sub>2</sub>) emissions based on a data panel covering 71 economies from 1996 to 2012. Specifically, we examine whether the level of income matters for the effect of green technology innovations. It is found that the impact of green technology innovations exists a single threshold effect regarding the income level. Specifically, green technology innovations do not significantly contribute to reducing CO<sub>2</sub> emissions for the economies whose income levels are below the threshold while the mitigation effect becomes significant for those whose income levels surpass the threshold. But the transition of regime occurs at an extremely high-income level. In addition, we find that the relationship between per capita CO<sub>2</sub> emissions and per capita GDP is inverted U-shaped, and urbanization level, industrial structure, trade openness, and energy consumption structure also significantly affect CO<sub>2</sub> emissions. Finally, this paper suggests that mechanism innovations should be implemented to reduce the diffusion cost of green technology in undeveloped economies.

## 1. Introduction

It is widely acknowledged that human activity such as burning coal and oil is one of the leading causes of global warming. Ever since the Industrial Revolution, the global economy has been evolving at a fast pace, and people's living conditions have been significantly improved, but improved productivity also brought severe air pollution worldwide. The World Energy Outlook 2017 cautions: "Despite their recent flattening, global energy-related CO<sub>2</sub> emissions increase slightly to 2040 in the New Policies Scenario. This outcome is far from enough to avoid severe impacts of climate change." Therefore, human activity is the genesis of global warming, and now humans are in urgent need of taking effective measures to protect the earth from climate disasters. Among various paths of climate change mitigation, the green technology (including renewable energy technology, energy efficiency technology, etc.) is expected to be a dominant factor that theoretically contributes to over 60% of targeted CO<sub>2</sub> reduction in the International Energy Agency's (IEA's) 450 Scenario (IEA, 2013). But in different countries or regions, the research development and diffusion of green technology are typically not at the same pace. Hence the actual impact of green technology innovations might depend on specific social or

economic circumstances (IEA, 2015). Thus, understanding the detailed relationship between human activity, green technology innovations, and CO<sub>2</sub> emissions helps to protect the environment that we depend on.

Since Grossman and Krueger (1991) first postulate the Environmental Kuznets Curve (EKC) hypothesis (which suggests an inverted U-shaped relationship between indicators of environmental pollutions and per capita income), a growing number of studies have devoted to investigating the factors affecting CO<sub>2</sub> emissions (Gill et al., 2018; Lean and Smyth, 2010; Liu et al., 2017; Perman and Stern, 2003; Stokey, 1998; Yang et al., 2015). Influencing factors such as prosperity, industrial structure, international trade, urbanization and energy structure have been discussed intensively. For instance, Yao et al. (2018) find that urbanization contributes to declines in China's CO<sub>2</sub> emissions. Munir and Ameer (2018) show that trade openness increases SO<sub>2</sub> emissions while urbanization reduces SO<sub>2</sub> emissions in Asian emerging economies. Sun et al. (2019) find that urbanization aggravates environmental pollution in China. Li et al. (2019b) reveal that the impact of manufacturing structural rationalization on CO<sub>2</sub> emission mitigation is subjected to the level of resource dependence and industrialization.

Recently, green technology innovations have grown up to be an important means of reducing CO<sub>2</sub> emissions all around the globe

\* Corresponding authors.

E-mail addresses: [kerrydu@xmu.edu.cn](mailto:kerrydu@xmu.edu.cn) (K. Du), [zhemingyan@snnu.edu.cn](mailto:zhemingyan@snnu.edu.cn) (Z. Yan).

(Weina et al., 2016; Nikzad and Sedigh, 2017). Although it is theoretically predicted that the higher the number of climate-related technologies the better for combating climate change, there are very few empirical evidences to support this (Su and Moaniba, 2017). Some previous studies suggest that the effect of green technology innovations on CO<sub>2</sub> emissions can be positive or negative under different conditions (Acemoglu et al., 2012; Jaffe et al., 2002), and can also be influenced by various factors, such as income and time. Braungardt et al. (2016) demonstrate that even though green innovations are generally considered as an essential element towards a green growth strategy, the impact on climate goals has been subjected to a long-running debate due to the existence of the rebound effect. Wang et al. (2012) find that energy technology patents do not play a significant role in reducing China's CO<sub>2</sub> emissions and energy patents with free-carbon technologies contribute to CO<sub>2</sub> emission reduction only in the eastern area of China. Weina et al. (2016) reveal that for Italia green innovations improve environmental productivity but not play a significant role in CO<sub>2</sub> emission reduction. Song et al. (2018) use the afforestation expanse from the environmental technology input as the proxy of green technology and explore its role in R&D efficiency and profit in manufacturing.

Understanding the real effect of green innovations in minimizing CO<sub>2</sub> emission deserves further study. Based on the existing studies, we pose two fundamental questions which need to be addressed. First, can green technology innovations effectively reduce CO<sub>2</sub> emissions? Second, are there some regime transitions for the effect of green technology innovations on CO<sub>2</sub> emissions under different income levels? As noted by Popp (2012), the using of green technology often entails an initial cost, which makes the poor economies unable to use advanced abatement technology and to achieve environmental goals.

This paper aims to empirically explore the above questions in depth using a new data set. Contributions of this paper are mainly twofold. First, the existing studies mainly focus on the impact of general technological advancement on CO<sub>2</sub> emissions. But few studies investigate the role of green technology innovations. This paper provides new evidence on the effect of green technology innovations on CO<sub>2</sub> emissions. Second, previous studies generally treat green technology innovations and income as general explanatory variables of CO<sub>2</sub> emissions, thus neglecting the interaction effect of income and green technology innovations on CO<sub>2</sub> emissions. Intuitively, the impact of green technology innovations might depend on the income level since using green technologies usually entail high costs. This paper is among the first to make income as a threshold to study the effects of green technology innovations on CO<sub>2</sub> emissions at different income levels.

The rest of the paper is organized as follows. In Section 2, we explain the econometric methodology. Section 3 details the data and results. Section 4 concludes the paper.

## 2. The model and econometric methodology

To investigate the effect of green technology innovations and income on CO<sub>2</sub> emissions, we consider the following reduced-form econometric model:

$$\text{Ln}(\text{Per\_CO}_2)_{it} = \text{Ln}(\text{Patent})_{it}\beta_1 + X'_{it}\gamma + u_i + \varepsilon_{it} \quad (1)$$

where  $\text{Ln}(\text{Per\_CO}_2)_{it}$  is the dependent variable defined as the logarithm per capita CO<sub>2</sub> emissions of economy  $i$  in year  $t$ .  $\text{Ln}(\text{Patent})_{it}$  is the core explanatory variable that denotes the logarithm of the number of green technology patents applied by economy  $i$  in year  $t$ .<sup>1</sup>  $X'_{it}$  represents a vector of control variables including per capita GDP, industrial structure, urbanization level, energy consumption structure, and trade openness, etc.  $u_i$  is the individual effects of economy  $i$ , and  $\varepsilon_{it}$  is the

<sup>1</sup> Note we actually use the logarithm transformation of (1 + patents) to avoid generating missing values when patents = 0.

random error.

To further investigate whether the effects of green technology innovations depend on the level of income, we include a dummy variable  $D$  which equals to 1 if the economy belongs to the high-income group.<sup>2</sup>

$$\text{Ln}(\text{Per\_CO}_2)_{it} = \text{Ln}(\text{Patent})_{it}\beta_1 + [\text{Ln}(\text{Patent})_{it} \times D]\beta_2 + X'_{it}\gamma + u_i + \varepsilon_{it} \quad (2)$$

By the introduction of  $D$ , we can easily compare the effects of green technology innovations between different income groups. However, Eq. (2) still has some disadvantages. Firstly, the sample separation criteria are exogenous. Secondly, economies are assigned to a specific group that is not changed throughout the given years; but with the development of economy, some less developed economies will enter the high-income group. To solve these problems, we use Hansen (1999) panel threshold model. The single threshold model is written as:

$$\begin{aligned} \text{Ln}(\text{Per\_CO}_2)_{it} &= [\text{Ln}(\text{Patent})_{it}I(q_{it} \leq \gamma)]\beta_1 + [\text{Ln}(\text{Patent})_{it}I(q_{it} > \gamma)]\beta_2 + X'_{it} \\ &\quad \gamma + u_i + \varepsilon_{it} \end{aligned} \quad (3)$$

where  $q_{it}$  is the threshold variable,  $\gamma$  is the threshold parameter that divides Eq. (2) into two regimes with coefficients of  $\beta_1$  and  $\beta_2$ . But generally, there may be  $K$  thresholds; and the model is written as:

$$\text{Ln}(\text{Per\_CO}_2)_{it} = [\text{Ln}(\text{Patent})_{it}I(q_{it})]\beta + X'_{it}\gamma + u_i + \varepsilon_{it} \quad (4)$$

where  $\beta$  is a  $K$ -dimensional vector,  $K$  is the number of thresholds,<sup>3</sup>  $I(\cdot)$  is a vector of indicator functions of which the  $k$ th component can be expressed as:

$$I_k(q_{it}) = \begin{cases} 1, & \text{if } c_{k-1} < q_{it} \leq c_k \\ 0, & \text{otherwise} \end{cases} \quad (5)$$

where  $k \in \{1, \dots, K+1\}$ ;  $c_0 = -\infty$ ,  $c_{K+1} = +\infty$  and  $c_1, \dots, c_K$  are  $K$  threshold parameters for estimation. The panel threshold model has been popularly applied to explore the nonlinear relationship between independent and dependent variables (Li et al., 2019a).

## 3. Data description and results

In this paper, we compile a balanced data panel covering 71 economies from 1996 to 2012. The economy list is provided in Table A1. The variables are constructed as follows.

### 3.1. Per capita CO<sub>2</sub> emission (denoted as Per\_CO<sub>2</sub>)

Following Ahmed et al. (2017) and Su and Moaniba (2017), we use per capita CO<sub>2</sub> emission as the proxy of CO<sub>2</sub> emission performance. The data of CO<sub>2</sub> emissions are collected from the World Bank.

### 3.2. Green patent counts (denoted as Patent)

Green technology innovation is an effective tool to address the conflict between economy and environment; it can effectively improve the energy efficiency which is vital to reduce CO<sub>2</sub> emissions (Braungardt et al., 2016). In this paper, we follow Su and Moaniba (2017) and Hasan and Tucci (2010) to utilize patent counts in environment-related technologies as the indicator of green technology innovations.<sup>4</sup> The data on green technology patent counts are collected

<sup>2</sup> We group the economies based on the World Bank's classification of economies (<http://data.worldbank.org/about/country-classifications/country-and-lending-groups>).

<sup>3</sup> The number of thresholds can be determined through the procedure developed in Hansen (1999).

<sup>4</sup> Another widely used indicators is expenditure on research and development (R&D). Expenditure on R&D is the input of innovation activities. It might not be in line with technology advancements which is generally taken as the out-come

from OECD statistics database (Yan et al., 2017). Referring to the methodology of OECD statistics, the patents labeled as “environment-related technologies” are counted and assigned to different countries according to bibliographic information on the inventor’s residential country.

### 3.3. Per capita GDP (denoted as *Per\_GDP*)

The studies of Environmental Kuznets Curve (Esteve and Tamarit, 2012; Tucker, 1995) highlight the non-linear relationship between per capita CO<sub>2</sub> emissions and per capita GDP. To test the EKC (Environmental Kuznets Curve) hypothesis, we include per capita GDP in our models as Du et al. (2012) have done. The data on real GDP and population are collected from PWT version 9.0 and the World Bank respectively.

### 3.4. Trade openness (denoted as *Ratio\_trade*)

Grossman and Krueger (1991) argue that the effect of trade openness on CO<sub>2</sub> emissions can be decomposed into scale effect, structure effect and technology effect respectively. Firstly, trade liberalization decreases friction, which will further enhance the scale of production, thus affecting the CO<sub>2</sub> emissions. Secondly, trade openness can affect industrial structure through specialization, which will further affect CO<sub>2</sub> emissions. Thirdly, technology can be transferred from technologically advanced economies to backward economies. It makes technology importer increases the efficiency of energy utilization, which may reduce CO<sub>2</sub> emissions. In this paper, we use the ratio of trade to GDP as the proxy of trade openness. The data are collected from the World Bank.

### 3.5. Urbanization level (denoted as *Ratio\_urban*)

Du et al. (2012) point out that the relationship between urbanization level and CO<sub>2</sub> emissions might be uncertain. Firstly, urbanization will inevitably facilitate the development of urban infrastructure, which may facilitate energy consumption and CO<sub>2</sub> emissions. Secondly, as the center of regional development, the urban area has a strong agglomeration effect, and the urban area can benefit from the scale effect in energy use, which may reduce CO<sub>2</sub> emissions. We use the share of urban population as the proxy of urbanization level. The data are directly collected from the World Bank.

### 3.6. Industrial structure (denoted as *Ratio\_ind*)

Industrial structure influences CO<sub>2</sub> emission mainly through two channels. Firstly, the change of industrial structure may affect the income growth of an economy, thus directly affecting the emission of CO<sub>2</sub>. Secondly, since the second industry is more energy-intensive and pollution-intensive, the industrial structural change may directly influence CO<sub>2</sub> emissions. Considering this effect, we use the output share of the industrial sector in the whole economy as the proxy of industrial structure. The data are directly obtained from CSMAR database.

### 3.7. Energy consumption structure (denoted as *Ratio\_renew*)

Different energy has different CO<sub>2</sub> emission coefficient.<sup>5</sup> Compared to renewable energy, fossil fuels emit more CO<sub>2</sub> given equivalent fuels

(footnote continued)  
of innovation activities.

<sup>5</sup> Generally speaking, coal products have the highest coefficient of CO<sub>2</sub> emissions in fossil energy followed by oil products and natural gas (Wang and Feng, 2017). In contrast to fossil energy, renewable energy such as wind and solar energy do not emit CO<sub>2</sub> directly.

and cause more damage to the environment (Capellan-Perez et al., 2014; Liu et al., 2017; Seow et al., 2016). Many economies are starting to adjust their energy consumption structures to reduce per unit GDP CO<sub>2</sub> emission (Kahia et al., 2016). Considering this effect, we use the share of renewable energy consumption in the total energy consumption as the proxy of energy consumption structure. The data are collected from World Bank.

### 3.8. Per capita output gap ratio (denoted as *Ratio\_gap*)

In this paper we use per capita output gap ratio which is defined as  $\text{Ratio\_gap}_{it} = \text{Per\_GDP}_{it} / \max\{\text{Per\_GDP}_{it}\}$  as the threshold variable. It reflects the income level of an economy relative to the observation with the highest income during the sample period. In addition, it also satisfies the stationarity condition of the transition variable.

Table 1 reports the descriptive statistics of variables by groups. It shows remarkable variations of economic development, CO<sub>2</sub> emissions, and green technology innovations and structural features between high-income economies and middle-income economies. For instance, the mean of GDP per capita in the high-income group is about four times as that in the middle-income group; the mean of green technology patents is five times more than that in the middle-income group; the mean of CO<sub>2</sub> emissions per capita is about two times as that in the middle-income group.

## 4. Estimation results and explanations

### 4.1. Estimation results of exogenous sample segment

We use the panel fixed effect model to estimate the Eqs. (1) and (2). The results are summarized in Table 2. According to Table 2, we can see that in all the models the Hausman test significantly rejects the null hypothesis, suggesting that there is correlation between regressors and the unobserved individual effects. Thus, using the fixed effects estimator to estimate the Eqs. (1) and (2) turns out to be reasonable. The result in Model I shows that the coefficient of  $\text{Ln}(\text{Patent})$  is estimated as  $-0.0134$ , insignificant at the 10% level. It indicates that overall, we do not find evidence supporting that green technology innovations can effectively curb CO<sub>2</sub> emissions. In Model II we consider the intersection term of green technology innovations and the group classification ( $\text{Ln}(\text{Patent}) \times D$ ). The result shows that the coefficient of  $\text{Ln}(\text{Patent})$  is estimated as  $0.0127$ , not significant even at the 10% level; but the coefficient of  $\text{Ln}(\text{Patent}) \times D$  is estimated as  $-0.0660$  and significant at the 1% level. Considering that there may be bidirectional causality between green technology innovations and CO<sub>2</sub> emissions, we replace  $\text{Ln}(\text{Patent})$  by the first-order lag term  $\text{Ln}(\text{Patent})_{-1}$  in Model III and Model IV. The results show a similar picture. In Model III, the coefficient of  $\text{Ln}(\text{Patent})_{-1}$  is estimated as  $-0.0129$ , insignificant at the 10% level. In Model IV, this coefficient becomes  $0.0156$ , not significant even at the 10% level while the coefficient of  $\text{Ln}(\text{Patent})_{-1} \times D$  is estimated as  $-0.0710$ , significant at the 1% level. It indicates that green technology innovations have negative effect on CO<sub>2</sub> emission only in the high-income group in which a 1% increase in green technology innovations would lead to a 0.0710% decrease in CO<sub>2</sub> emissions.

The results above are consistent with our intuition. We explain the results from two aspects. (1) From the perspective of the market, the application of green technology is on a cost-benefit basis that depends on an economy’s prosperity status to a certain degree. Due to the relatively poor productivity and low marketization level in the undeveloped economies, applying green technology in practice would lead to high manufacturing cost. Consequently, there might be some green technology innovations in the underdeveloped economies, but the cost of diffusion is typically unaffordable for local firms or residents. (2) From the perspective of government, available governmental resources vary significantly across economies with different income levels, which makes the motivation of developing green technology

**Table 1**  
Summary statistics of variables by groups.

Group	Variable	Unit	Mean	SD	Min	Median	Max
M	Patent	–	101.764	472.292	0	5.29	5581.92
	Per_CO <sub>2</sub>	Tons/person	3.861	2.873	0.294	3.63	12.811
	Per_GDP	2011US\$/person	8154.931	4702.145	1427.32	7394.929	24,065.33
	Ratio_trade	%	77.125	39.66	0.027	68.025	220.407
	Ratio_urban	%	57.115	18.854	18.297	56.778	91.295
	Ratio_ind	%	19.168	8.375	0.65	17.795	58.98
	Ratio_renew	%	9.975	13.316	0	4.424	64.817
H	Patent	–	626.07	1692.144	0	60.87	11,292.07
	Per_CO <sub>2</sub>	Tons/person	9.227	4.12	2.678	8.508	25.221
	Per_GDP	2011US\$/person	31,078.91	12,271.3	8617.783	31,299.77	84,270.29
	Ratio_trade	%	103.255	76.912	18.756	80.845	449.993
	Ratio_urban	%	76.582	12.303	49.856	77.889	100
	Ratio_ind	%	16.729	5.691	1.52	16.91	31.37
	Ratio_renew	%	14.974	16.933	0	10.478	89.725

Note: M and H represent middle-income and high-income groups, respectively.

**Table 2**  
Estimation results of panel data model with exogenous sample segment.

	Model I	Model II	Model III	Model IVs
Ln(Patent)	-0.0134 (0.0143)	0.0127 (0.0153)		
Ln(Patent) × D		-0.0660** (0.0275)		
Ln(Patent) <sub>-1</sub>			-0.0129 (0.0130)	0.0156 (0.0136)
Ln(Patent) <sub>-1</sub> × D				-0.0710*** (0.0262)
Ln(Per_GDP)	1.6228*** (0.5781)	1.1502** (0.5500)	1.5959*** (0.5636)	1.1261** (0.5360)
[Ln(Per_GDP)] <sup>2</sup>	-0.0807** (0.0310)	-0.0549* (0.0292)	-0.0791** (0.0303)	-0.0536* (0.0285)
Ratio_trade	-0.0006 (0.0006)	-0.0005 (0.0006)	-0.0007 (0.0006)	-0.0005 (0.0006)
Ratio_urban	0.0251*** (0.0077)	0.0235*** (0.0073)	0.0249*** (0.0080)	0.0233*** (0.0075)
Ratio_ind	0.0065* (0.0037)	0.0049 (0.0039)	0.0064* (0.0035)	0.0046 (0.0037)
Ratio_renew	-0.0114*** (0.0042)	-0.0112*** (0.0041)	-0.0112*** (0.0041)	-0.0109*** (0.0039)
Constant	-8.0190*** (2.7561)	-5.6953** (2.6807)	-7.9022*** (2.7164)	-5.5749** (2.6261)
Hausman test	54.47*** {0.000}	71.84*** {0.000}	55.32*** {0.000}	74.92*** {0.000}
No. of observations	1207	1207	1136	1136
No. of id	71	71	71	71

Note: Robust standard errors in parentheses; P-value in brace.

- \*\*\* p < 0.01.
- \*\* p < 0.05.
- \* p < 0.1.

inconsistent among groups. For the group of undeveloped economies, governments concentrate more on developing the economy to improve people's living standards. Due to lack of enough material foundation, it is costly for them to promote green technology innovations into real applications. Thus, the impacts of green technology innovations on CO<sub>2</sub> emission are not significant in undeveloped economies. On the contrary, for developed economies, people's desire for a better living environment will be stronger, and governments are more capable of promoting green technology innovations into real application in society. Therefore, green technology innovations play a significant role in mitigating CO<sub>2</sub> emission in the high-income group.

#### 4.2. Estimation results with endogenous thresholds

The above analysis is based on the exogenous sample segment. In this subsection, we use panel threshold models to further investigate

**Table 3**  
Estimation results of panel data model with endogenous thresholds.

	Model V	Model VI
<i>Threshold estimates</i>		
c	0.4117	0.4117
95% confidence interval	[0.4049,0.4139]	[0.4045,0.4122]
Ln(Patent) × I(Ratio_gap ≤ c)	-0.0098 (0.0071)	
Ln(Patent) × I(Ratio_gap > c)	-0.0283*** (0.0074)	
Ln(Patent) <sub>-1</sub> × I(Ratio_gap ≤ c)		-0.0099 (0.0072)
Ln(Patent) <sub>-1</sub> × I(Ratio_gap > c)		-0.0277*** (0.0076)
Ln(Per_GDP)	1.2629*** (0.1777)	1.2800*** (0.1794)
[Ln(Per_GDP)] <sup>2</sup>	-0.0610*** (0.0099)	-0.0618*** (0.0100)
Ratio_trade	-0.0005* (0.0002)	-0.0005** (0.0002)
Ratio_urban	0.0267*** (0.0021)	0.0270*** (0.0022)
Ratio_ind	0.0048*** (0.0015)	0.0048*** (0.0016)
Ratio_renew	-0.0119*** (0.0013)	-0.0118*** (0.0013)
Constant	-6.4770*** (0.8108)	-6.559*** (0.8167)
<i>Likelihood ratio test</i>		
H <sub>0</sub> : Linearity	47.75* {0.0760}	43.87* {0.0680}
H <sub>0</sub> : One threshold	22.55 {0.4320}	25.61 {0.2680}
No. of observations	1207	1136
No. of id	71	71

Note:

- (1) Standard errors in parentheses; P-value in brace.
- (2) The bootstrapping times for the likelihood ratio test is 500.

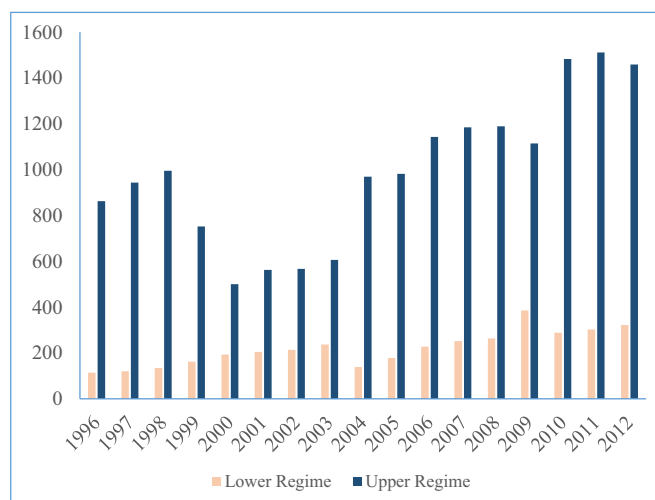
- \*\*\* p < 0.01.
- \*\* p < 0.05.
- \* p < 0.1.

the interaction effects of green technology innovations and income on CO<sub>2</sub> emissions. The results are presented in Table 3. In Model V and Model VI, the results of the likelihood test show that the linearity hypothesis is rejected at the 10% significant level, but the single threshold hypothesis cannot be rejected. It means that there are regime transitions for the effect of green technology innovations on CO<sub>2</sub> emissions which depends on the income level. In other words, when per capita income reaches a certain level, the effects of green technology innovations on CO<sub>2</sub> emissions would change. The threshold parameter is estimated at 0.4117 (it is equivalent to 34,694.0782011 US dollars per

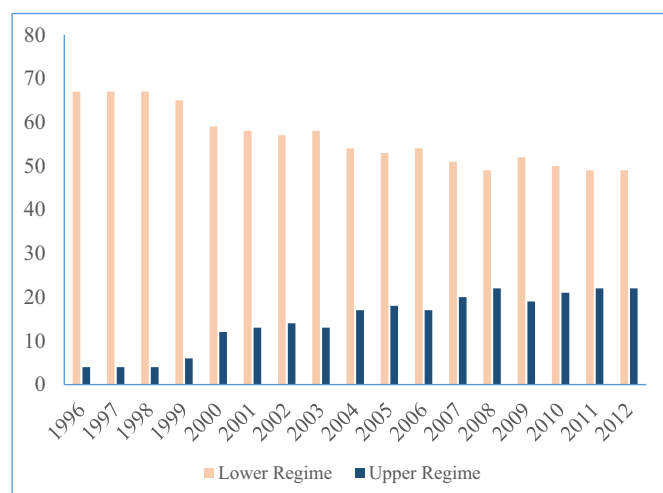


**Table 4**  
The group classification based on the estimated threshold parameter.

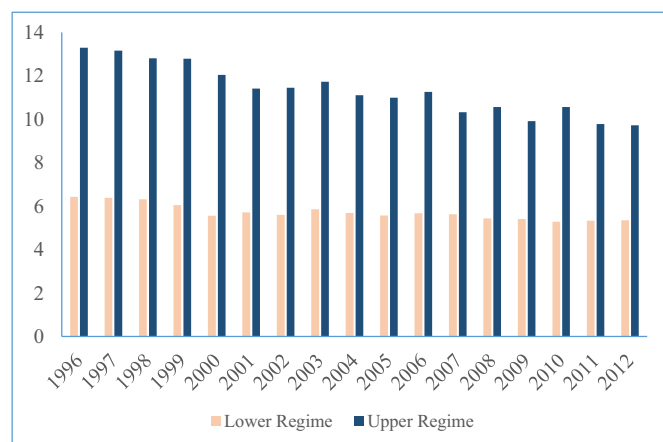
Classification	Economy			
Upper regime	Switzerland	Luxembourg	United States of America (USA)	
Transition from lower regime to upper regime	Norway	Austria	Belgium	Canada
	Australia	Denmark	Finland	France
	Germany	Denmark	Finland	France
	Great Britain (United Kingdom)	Italy	Hong Kong	Ireland
	Iceland	Singapore	Japan	Netherlands
Lower regime	Saudi Arabia	Singapore	Sweden	
	Argentina	Armenia	Azerbaijan	Bulgaria
	Belarus	Brazil	Chile	China
	Colombia	Cyprus	Czech Republic	Spain
	Estonia	Georgia	Greece	Croatia
	Hungary	Indonesia	India	Iran
	Iraq	Israel	Kyrgyzstan	South Korea
	Lebanon	Sri Lanka	Lithuania	Latvia
	Morocco	Moldova	Mexico	Malaysia
	New Zealand	Pakistan	The Philippines	Poland
	Portugal	Romania	Russian Federation	Slovakia
	Slovenia	Thailand	Tajikistan	Turkmenistan
	Turkey	Ukraine	Uzbekistan	Venezuela
	South Africa			



**Fig. 3.** Green technology patents by groups over times. (For interpretation of the references to colour in this figure legend, the reader is referred to the web version of this article.).  
Note: Average values by group are presented.



**Fig. 1.** Number of economies by groups over times.



**Fig. 2.** Carbon dioxide emissions by groups over times.  
Note: Average values by group are presented (Unit: Tons/person).

capita). Table 4 shows the sample grouping results based on the estimated threshold parameter. We find that only four most developed economies (Norway, Switzerland, Luxembourg and the United States) lie in the upper regime; 49 developing economies such as China, India, and Brazil lie in the lower regime; the other 18 economies such as Britain and Japan successfully realized the transformation from the lower regime to the upper regime during the research period. Fig. 1 shows the number of observations in upper and lower regimes over times. It can be seen that before 1999 the number of economies in the upper regime is staying at 4, and after that, the number increases continuously and eventually becomes stable around 20. Fig. 2 presents a comparison of CO<sub>2</sub> emissions between different regimes. It shows that CO<sub>2</sub> emissions declined substantially in the upper-regime group while the lower-regime group did not evidence a declining trend obviously. Fig. 3 compares green technology innovations between the two groups. We can find that there are large gaps of green technology innovations between the lower and upper regimes; furthermore, the gaps are enlarging over times.

For Model VI, when *Ratio\_gap* is < 0.4117, the estimated coefficient of  $\ln(\text{Patent})$  is  $-0.0099$ , insignificant at the 1% level; When *Ratio\_gap* is > 0.4117, the coefficient of  $\ln(\text{Patent})$  is estimated as  $-0.0277$ , significant at the 1% level. The results reported in Model VI show a similar picture. Thus it is a robust result that green technology innovations contribute to reducing CO<sub>2</sub> emissions only when the income level of the economy surpasses a certain level. In other words, green technology innovations do not play a significant role in CO<sub>2</sub> emission reduction for underdeveloped economies.

With regard to the control variables, we obtain some conclusions from Table 3. The coefficients of  $\ln(\text{Per\_GDP})$  are all significantly positive, and the coefficients of  $[\ln(\text{Per\_GDP})]^2$  are all significantly negative, indicating that the Environmental Kuznets Curve hypothesis is documented. That is to say, the inverted U-shaped relationship between per capita CO<sub>2</sub> emissions and per capita GDP is supported. The coefficient of *Ratio\_renew* is negative and significant at the 1% level, indicating that the increasing share of renewable energy consumption would lead decline in CO<sub>2</sub> emissions. Similarly, the coefficients of *Ratio\_trade* are negative and significant even at the 5% level, suggesting that there is a negative relationship between foreign trade and CO<sub>2</sub> emissions. In other words, we find a contradictory evidence of “pollution haven hypothesis”. The coefficients of *Ratio\_urban* and *Ratio\_ind* are all positive, indicating that the increases in the urbanization level and the output share of the industrial sector have positive effects on CO<sub>2</sub>

emissions. The results are similar to the findings of Li and Lin (2015).

## 5. Conclusion and policy implications

This paper explores the heterogeneous impact of green technology innovations on CO<sub>2</sub> emissions by using the panel threshold model proposed by Hansen (1999). Our empirical study provides ample evidence that income levels drive the non-linear nexus between green technology innovations and CO<sub>2</sub> emissions. We find that the effect of green technology innovations exists a single threshold effect with regard to the income level. To be specific, the effect of green technology innovations on reducing CO<sub>2</sub> emissions is more significant for the economies whose income level surpass 34,694.078 US dollars in the 2011 price level. It means that the transition of regime occurs at an extremely high-income level.

The above results have some policy implications for climate change mitigation. Firstly, the world is in urgent need of fostering developing economies' green innovation capacities, given that combating climate change is the mutual task for all economies, and that green technology innovation can reduce CO<sub>2</sub> emissions while stimulating economic growth. Secondly, green technology innovations are still playing a vital role in climate change mitigation (Zhang et al., 2016), although they only take effect when the economy reaches a high-income level. Zhang et al. (2018) document that scientific research funds play an important role in stimulating technological progress. Thus, the government of the developed economies should allocate more resources in R&D and encourage enterprises to make green technology innovations. Thirdly, it is necessary to construct a new framework regarding worldwide diffusion and application of green technology. Since the green technology is typically expensive for individuals in low-income economies, some mechanism innovations in intellectual property, green finance, and governmental support should be initiated for accelerating diffusion and application of green technology. Additionally, the low-income economies should promote the application of green management which can improve resource utilization efficiency given the production technology (Raharjo, 2019). Fourthly, while only three economies were on the right side of the inverted U-shaped curve from 1995 to 2012, another 19 economies successfully reached the turning point for CO<sub>2</sub> emissions; thus, governments of the lower regime should vigorously seek paths of green growth and raise the national income level to transform at an earlier date.

Finally, it is worth pointing out that this paper attempts to provide new evidence on the heterogeneous impact of green technology innovations on CO<sub>2</sub> emissions whereas measuring green technology innovations is a challenging issue. Due to data availability, we employ the patent counts in environment-related technology as the proxy of the green technology innovations which has some potential limitations. Although the patent-based indicator has been widely employed in the existing literature, the potential limitations should be also duly noted. First, it can reflect technological development, but cannot represent the situation of technology adoption.<sup>6</sup> Second, patent counts simply add up the patents which might not neglect various values of different innovations.

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<sup>6</sup> In view of this, some recent studies attempt to account for environmental technology progress based on the production-theory framework. Examples of such studies include Shao et al. (2016), Yu et al. (2017), Song and Wang (2018), and Jin et al. (2019). However, in this context some particular assumptions are required for estimation. Consequently, different methods generally lead to greatly various estimates of environmental technology progress.

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Table A1

List of economies and the group classification based on World Bank.

Classification	Economy				
High income D = 1	Australia	Austria	Belgium	Canada	
	Switzerland	Chile	Cyprus	Czech Republic	
	Germany	Denmark	Spain	Estonia	
	Finland	France	Great Britain (United Kingdom)	Hungary	
	Greece	Hong Kong	Croatia	Italy	
	Ireland	Iceland	Israel	Luxembourg	
	Japan	South Korea	Lithuania	New Zealand	
	Latvia	Netherlands	Norway	Singapore	
	Poland	Portugal	Saudi Arabia		
	Slovakia	Slovenia	Sweden		
	United States of America (USA)				
	Middle income D = 0	Argentina	Armenia	Azerbaijan	Bulgaria
		Belarus	Brazil	China	Colombia
		Georgia	Indonesia	India	Iran
Iraq		Kyrgyzstan	Lebanon	Sri Lanka	
Morocco		Moldova	Mexico	Malaysia	
Pakistan		The Philippines	Romania	Russian Federation	
Thailand		Tajikistan	Turkmenistan	Turkey	
Ukraine		Uzbekistan	Venezuela	South Africa	

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